# Armed Services Technical Information Agenc

NOTICE: WHEN GOVERNMENT OR OTHER DRAWINGS, SPECIFICATIONS OF OTHER DATA ARE DISO FOR ANY PURPOSE OTHER THAN IN CONNECTION WITH A DEFINITE IT RELATED GOVERNMENT PROCUREMENT OF HRATION, THE U.S. GOVERNMENT THEPERY INCURS NO RESPONSIBILITY, NOR ANY OBLIGATION WHATSONVER; AND THE FACT TEXT THE GOVERNMENT DAY HAVE FORMULATED, FURNISHED, GRIN ANY VAN SUPPLIED THE SAID DRAWINGS, SPECIFICATIONS, OR OTHER DATA IS FOR TO PEREGUOSE BY DISPLANT ON OTHER AS IN ANY MAIN IN LICENSIBLE OF PERECUENCE OR ANY OTHER PERSON OR COMPORATION, OR CONVEXES ANY PIGHTS OF PERSON TO I ANUSACTURE, U.S. CH SELL ANY PATINITED INVENTION THAT MAY IN ANY WAY LE BELATED THERETO.

ENT SERVICE CENTER

AD047524

PRS REPORT NO. 984

A METHOD FOR SYMPHESIS

OF FACTOR ANALYSIS STUDIES

bу

LEDYARD R TUCKER Principal Technician

EDUCATIONAL TESTING SERVICE

PRINCETON, M. J.

FIRAL REPORT

31 March 1951

DEPARTMENT OF THE APMY
Project No. 29530100 Subtask 75
RESEARCH CONTRACT NO. DA-1-9-033 CSA- 116
PERSONNEL RESEAPOT SECTION PR 3075

PRS reports are primarily technical. While conclusions affecting military policy or operations may appear in them, they are not intended as a basis for official action. Findings and conclusions contained in PRS reports are intended to guide the conduct of further research. When research findings suggest recommendations for administration action, such recommendations are made separately to the appropriate military agency.

This report is of primary interest to research workers who have occasion to use factor analysis techniques. Since factor analysis is a basic research tool, the developments described in this report will be of interest to research workers in personnel management as well as in other fields.

The technique of factor analysis is a widely used and important research tool aimed at giving a better understanding of the underlying abilities that personnel classification tests measure. In the area of test construction and administration, the method of factor analysis can be used to answer the question: How many traits, or abilities -- or for convenience, factors -- are measured by a given test or set of tests? Such information is used in improving the effectiveness of tests and test batteries. To date, the technique of factor analysis has been developed to produce this information for the tests used in only one study at a time. The problem arises, then, of comparing the factors isolated in different experimental studies to ansver the further question: Are the factors isolated in each of several studies identical or different? To answer such questions, it has been necessary to incorporate the tests from several studies into a larger, over-all study on a representative sample of people and then apply factorial analysis techniques. The difficulties and cost of doing this have discouraged its being done.

It is, therefore, desirable to have a technique for making more immediate comparisons of factors. The present investigation, primarily methodological in nature, is concerned with a technique that will allow comparison of factors when certain conditions are met and will also enable adjustment of the factors obtained in separate studies so as to synthesize the findings. Computational details of the technique are described, and several numerical examples are presented to show the practicability of the technique.

## CONTENTS

প্রয়োক তিনা এন বাশিক স্বয়ন্ত্রা জন্ম তথ্যসূত্র বুদ্ধান বিশ্বনার বিশ্বনার স্থানিক স্থানিক বিশ্বনার বিশ্বনার স

Section	on.		Page
ı.	Int	production	1
n.	Res	ults of a Search of the Literature	· 5
	۸.	Controversey on Invariance of Factorial Results When Several Studies are Analyzed in Accordance	5
		with Thurstone's Multiple Factor Analysis Theory	5
	В.	Solutions to Similar or Related Problems	9
III.	Ger	eral Conditions and Assumptions	11
	A.	Assumption of Thurstons's General Theory of Factor Analysis	11
	в.	Effects of the Groups Tested	11
	C.	Defined Restrictions	12
	D.	Synthesis of Studies Tvo at a Time	15
-	В.	Some Tests Common to Both Batteries, Other Tests in Either Battery	15
ĬĀ	Cons	gruence, of Fector Matrices	18
	A.	Definition of Congruence	2 <b>9</b> 5
	в.	The Non-Congruent Spaces	20
	c.	Meaning of an Observed Congruence	20
Ref	erenc	03	<b>30</b>
APP	EKNI	A Theory of Multiple Factor Analysis of Covariance	32
	1.	General Factorial Equation for Covariances	32
	2.	Transformation of Factors	33
	3.	Transformation of Tests	<b>3</b> 6
	4.	Effects of Between-Group Differences	37
APPET	IDIX .	B Details of Minimum Solution for Index of Congruence	40
APPE	NDIX	Computing Procedure for Synthesis of Factor Analysis Studies	65
	1.	Congruent Factor Computations	66
	2.	Retation of Axes in the Congruent Space	89
	3.	Determination of Hon-Congruent Axes	94
	4.	Determination of Latent Roots and Vectors	102
	5.	Notes on Matrix Computations	116

A ministrations

# LIST OF TABLE

Table		Page
1	Rotated Factor Loadings	3-4
2 :	Adjusted Test Standard Deviations	16
3	Loadings on Reference Factors for Overlap Tests with Adjusted Standard Deviations	17
<b>4</b>	Loadings of Overlap Tests on Congruent Factors	24
5	Transformation to Congruent Factors	25
6	Direction Cosines of Mon-Congruent Factors	26
7	Loadings of Overlap Tests on Rotated Congruent Factors	27
8	Transformation to Rotated Congruent Factors from Reference Factors	<b>28</b> -
9	Loadings of Non-Overlap Tests on Rotated Congruent Factors	29
10	Matrices (F' Fina and (F' Fina)	54
11	Latent Vectors and Roots for Overlap Tests	55
12	Factor Loadings on Principal Axes	56
13	The Matrix (Final JPB)	57
14	The Matrices G, HA, HB	58
15	Latent Vectors and Roots for Matrices HA and HB	5 <del>9</del>
16	Matrices 7mrA, 7mrB, c	60
17	Summation Vectors for First Rotation in Congruent Space	61
18	Transformation Matrices for First Rotation in Congruent Space	62
<b>19</b> .	Factor Losdings for First Rotation in Congruent Space	63
20	Transformation Matrices for Second Rotation in Congruent Space	6'3
21-34	(Illustrative Example for Congruent Factor Computations, Appendix C, Section 1.)	76- 8 <b>8</b>
35-36	(Illustrative Example for Rotation of Axes in the Congruent Space, Appendix C, Section 2.)	92- 93
37-38	(Illustrative Example for Determination of Non- Congruent Axes, Appendix C, Section 3.)	100-101
39-44	(Illustrative Example for Determination of Latent Roots and Vectors, Appendix C, Section 4.)	110-115

# A NOTIFICE FOR SYNTHESIS OF FACTOR ANALYSIS STEDIES\*

Ledyard R Tucker

# I. Introduction

After several factor analysis studies in a particular domain have been completed, one of the major questions that occurs is how to synthesize results from these studies. Often there are a number of identical tests in the several batteries employed in the studies. It is hoped that these common tests will aid in a more firm synthesis of the studies. The problem has been how to make use of this identity of tests. Claims have been reported in the literature that rotation to simple structure will yield invariance of factors, and it has been held that the common tests could assist in identifying corresponding factors in the two studies. However, some difficulty has been encountered in this approach (as has also been reported in the literature); consequently, the need for devising a wore definite method has become clear. In particular. it has been hoped that by some new method two studies which have overlapping tests and have been factored in accordance with L. L. Thurstone's general theories to loadings on reference axes could be separately retated into congruence and then jointly rotated to simple structure.

An illustration of the problem under attack is provided by two studies conducted by the Personnel Research Section of the Personnel Research and Procedures Branch, Office of the Adjutant General, Department of the Army in cooperation with representatives of the other Armed Services. One study

\*The author is indebted to a number of staff members in the Department of Statistical Analysis, Educational Testing Service, who have worked on development of the material of this report. Deserving of special Lentlon are Mrs. Gertrude Diederich who assisted in the analysis of the major example and Miss Angela E. Nolan who assisted in the preparation of the computing directions. Portions of the manuscript were read by Dr. William G. Hollenkopf and Dr. Frederic M. Lord, both of the Educational Testing Service Research Department, and by Miss Henrietta L. Callaguer of the Department of Statistical Analysis.

involved Army and Havy classification tests\* which were given to a group of Haval Recruits after which a factor analysis was performed including rotation of axes to an orthogonal simple structure given in Table 1. This study will be called Study A. The other study, which will be called Study B, involved Army, Havy, and Air Force tests given to some Airmen and some Soldiers. A factor analysis again was performed with rotation to the orthogonal simple structure also given in Table 1. Ten of the variables were common to the two studies as indicated by the cross references included in the two tables. Some difficulty has experienced in identification of the same factors in the two studies. These two tables will be called matrices F<sub>jmA</sub> and F<sub>jmB</sub>. It is the purpose of the present investigation to develop a means for finding a factor space common between such studies as determined by similarity of factor loadings of tests common to the studies. This may be expressed mathematically as attempting to discover transformation matrices T<sub>mrA</sub> and T<sub>mrB</sub> such that when

(I.14

(I.1B

the differences between matrices  $F_{\underline{irA}}$  and  $F_{\underline{irB}}$  are negligible for tests overlapping the two studies. The subscripts  $\underline{m}$  and  $\underline{\underline{H}}$  have been used to designate the reference factors in the two studies valle  $\underline{\underline{r}}$  is used to designate what we will term congruent factors. A and  $\underline{\underline{B}}$  are used to designate the studies. The number of congruent ractors may be less than the number of factors in either study.

Once a set of congruent factors have been determined, further rotation of axes in both studies jointly within the space of the congruent factors can be accomplished to rotated congruent factors. Whenever the congruence has been established as of sufficient strength, these steps should facilitate across study comparisons.

<sup>\*</sup>TRS Report 778. Comparison of Army and Mavy Classification Tests. 29 April 1949.

### Nable 1

# ROTATED FACTOR LOADINGS

# STUDY A

# Matrix F.

Test Code	Numbers		-	Loading	•			
•				(Decimal	points	are or	itted.	
Study A	Study B	-	<u>i</u>	_11_	111	17	Y	<u></u>
, 9	26		83	0/4	01	01	08	-01
18	1		59	21	09	-01	07	-01
. 1	••	•	87	06	-02	23	oi.	19
5	23		57	-09	01	53	27	21
12	58	•	58	07	16	35	11	27
14	55		27	09	03	44	53	35
13	32		63	10	03	27	46	27
14	34		60	04	CO	10	30	80
15	33		40	00	01	37	56	00
11	27		78	05	41	33	01	00 .
3	••		70	16	38	26	00	-08
10	• •		72	15	46	10	-02	11
	••		62	22	46	01	09.	04
2 6	••		59	45	22	09	-09	10
7	• •		45	63	00	oí	02	00
ė.	• •		33	80	13	05	03	08
16	**		39	59	08	18	-01	03
17	30		52	50	10	03	· 03	00

<sup>\*</sup>These rotated factors will be used as reference factors in the following analysis and will be so labeled.

# Table 1 (Continued)

# NOTATED PACTOR LOADINGS

### STUDY

Test Code	. Numbers					Loadi	ags on	Notated :	Pactors		,	•	
<del>47 74 15 4.</del>			•										
Study A	Study B	1	11	III	IA	(Decri	AI bon	eta eze d VII	AIII Erstear).		T	n	XII.
							-					<del></del>	
18	1	451	-075	<b>303</b> 604	946 246	175	454	080	Oh1	063	-010	091 048	116
••	2 3	30 <b>4</b> 230	-300 061	630	430	-007 249	085	-077 -102	072 154	015 118	-187	102	125
••	ĭ	273	067	330	236	236	061	-023	244	225	-036	-015	015 -085
**	5	-004	320	-611	278	OHA	035	-066	-167	-049	-1.54	-025	Ok9
••	5	192 -040	-142	748	137	068	234	-145	165	-075	-005	108	186
••	7	-040	-505	-145	-215	-164	-154	070	-065	132	-088	-035	-298
••	8	192 690	-124	849	180	143	-029	-0 <b>13</b>	225	126	-088	145	-061
••	9	690	-043.	109	274	166	230 418	-00%	177	115	-052	285	088
••	10 11	396 290	-033 -038	070 134	257 265	074 329	500	-033	116 151	220 246	120 085	390 390	025 045
••	72	274	017	694	156	181	594	-000	078	135	039	256	-043
	13		025	079	370	229	110	-054 -054	257	173	O¥9	216	-003
••	14	590 668	-Oko	125	198	160	510	-035	110	120	107	243	075
••	15 16	306	646	100	578	147	151	-015	264	237	017	243 283	059
••		138	041	-015	455	290	104	-034	220	313	132 038	206	070
••	17	227	070	020	723	098	175	-014	124	149	038	277	019
••	18	078	014	-011	701	195 484	016	-009	043	238	089	288	075
••	19	221	050	109	167		174	-020	025	291	034	212	-010
••	20 21	225 157	055 -019	049 132	129 221	327 370	335 205	-05 <b>5</b> -057	152 105	139	-105	220 466	025
1	22	050	-073	020	615	228	008	014	135	355 380	-006	165	164
Š	23	182	043	049	532	265	151	-027	188	447	092	505	041
	24	492	-078	117	160	041	364	102	140	110	000	281	002
••		567	030	057	330	130	245	-056	284	183	-062	335	~069
9	5 <u>6</u> 52	550	~022	266	530	154	370	-031	245	174	-060	335 274	-037
11	27 28	385	-065	107	292	090	481	010	154	296	036	466 466	088
12	28	231	-00,4	124	263	359	228	009	680	446	-080	466	050
16,17	29	224	-042	003	495	180	250	-007	500	435	086	129	-056
	30 31	315 100	001 -007	099 095	120 207	341 366	550	-002 -071	065 21: <b>4</b>	180 001	-008 090	129 215	-001 -050
13	32	320	052	061	588	186	293 236	016	180	366	033	508	015
15	33	073	105	-014	765	-007	153	-048	147	230	-000	173	-015
ī¥	34	147	-045	Ohl	335	035	172	-055	587	171	055	270	-010
••		ooi	-057	320	178	CHA	049	-055 .064	617	033	-054	150	025
••	35 36	221	113	-005	-076	151	140	-407	-091	160	007	-198	-074
**	37	-1)42	599	-150	514	-099	-092	020	-115	093	-220	රයිට	-065
••	58	032	584	0.13	123	0.3	015	015	008	150	200	c84	075
••	59 10	-007	557	-011	410	-016	-035	159	-024	-025	-099	095	-102
••	40	134	916	227	154	90)	105	262	1164	220	-050	023 -194	316 -109
••	41 42	055 014	-657 -668	-070	-170 -217	-050 -017	-105 204	092 050	-075 002	-127 -095	-09 <del>4</del>	-194	224
::	43	070	-003 055	113	307	130	-015	439	222	-028	099	112	-050
••	44	-101	-354	-123	-303	058	-235	-136	-275	-184	164	154	-264

# II. Results of a Search of the Literature

As the study of the literature progressed it became apparent that publications to date could be classified under two headings. First, there has been some controversy on invariance of results from factorial studies. Second, several problems similar or related to the one we are studying have been treated and solutions have been found. Evaluation of the material presented in the literature became guided, as more material was covered, by a greater precision in definition of the general nature of the attack we were going to make on the solution of the problem.

A. Controversy on Invariance of Factorial Results When Several Studies
Are Analyzed in Accordance with Thurstone's Multiple Factor Analysis
Theory.

Thurstone has suggested (25, 26, 28) that rotation of axes in factor analyses to simple structure will yield invariant results for common factors under broad conditions of change from one study to another. Several limiting conditions in which invariance could not be expected were also noted. Specific changes for which invariance was claimed were:

- 1. The pattery of tests could be altered by addition of other variables or by deletion of variables, provided that:
  - (a) The new variables did not have loadings on a factor specific to one of the original variables, in which case this previous specific factor would become a new or mor factor; however, the previous common factors were invariant.
  - (b) The new variables did not have additional common factors among themselves, in which case a new common factor or factors would be added; however, the previous common factors were invariant.

(c) The variables deleted did not eliminate one of the previous common factors; however, the other common factors remained invariant.

(Thurstone was cognizant of the fact that deletion of tests could leave the configuration of vectors such that the simple structure was so indeterminate that the rotation of axes would not result in the same factors. Also recognized was that addition of new tests could help determine a simple structure previously indeterminate.)

- 2. The battery of tests could be administered to different groups of individuals, provided that:
  - (a) The groups were sufficiently similar so that the psychological nature of tasks involved in the tests did not change.
  - (b) Partial special selection had not occurred between groups on two or more variables. In this case a new "incidental" common factor would be added, but the simple structure would remain for the common factors.
  - (c) Complete special selection had not occurred on one or wore variables.

(Thurstone, in this area of change, considered it permissible for the factors to change in correlation and for the exact values of the non-vanishing factor loadings to change. The factor loading changes, however, are approximately by a constant of proportionality for each factor under theoretical conditions of selection.)

Meyer (16) has experimentally demonstrated factorial invariance when tests are deleted from a battery. He dealt explicitly with cases in which the test deletions left the simple structure indeterminate. In one of three small batteries formed by a selection of the tests used by the Thurstones in a study at the eighth grade level (27), one factor which should have been present could not be identified after separate analysis of the small battery and rotation to what appeared to be its simple structure. Meyer used Thurstone's correlations so that only invariance under test selection was being tested. It seems, however, that data from

Meyer's third battery as compared with Thurstone's results for a larger battery would be ideal for a tryout of any methods developed in the present project.

Godfrey Thomson (23), applying Pearson's and Aitken's formulation of effects of selection, has pointed out that universate and multivariate selection affected factorial solutions. Thurstone's analysis of the results of these effects has been previously noted.

Cyril Burt (1) has used the correlation between two sets of factor loadings as an indication of agreement. R. B. Cattell (7) called this a "shape correlation coefficient." William Stephenson (22) has used the rank correlation coefficient in his Q-technique as a measure of similarit of profiles. Cattell (7) also suggested a coefficient of pattern similarity:

$$r_{p} = \frac{2k - \Sigma d^{2}}{2k + \Sigma d^{2}}$$

in which k is the median for  $\chi^2$  on a sample of size n, the number of tests involved. (Each d would be the difference in factor loadings for one test in two analyses.) Cattell was particularly interested in similarity of profiles for two individuals and the d's would be the differences in their scores; but he suggested that the coefficient could be applied to similarity of factor loadings of tests in two studies.

Two other suggestions by Cattell were reviewed. In one (2) he suggested a different method for determining rotation of axes than the principle of simple structure. His suggestion was to wake two studies, using the came tests in both studies, but so operating on the groups of subjects examined that the variability of the tests attributable to one psychological function would differ between groups. Estation would be guided to that position where leadings on all factors but one would be equal between two studies and loadings on the one factor would be proportional between studies. Mathematics for accomplishing the desired result were presented only for the two dimensional case with no minor

inconsistencies arising from such sources as sampling errors. In another article (4) Cattell used bivariate frequency counts and probabilities to support matching of factors after two studies had been rotated to simple structure.

H. S. Reyburn with J. G. Taylor in 1945 (18) and with M. J. Reath in 1949 (19) has criticized simple structure as a basis for rotation of axes in factor analysis. Young and Householder (29) suggested pivoting rotations in successive studies on particular tests by which they hoped a large body of knowledge about relations could be built.

Points of particular interest in the foregoing review of literature on factorial invariance of simple structure solutions to the present project are:

- 1. Disagreement exists as to whether or not to expect simple structure factorial solutions to be invariant.
- 2. Several indices for identity of factors in two studies have been put forward. Of most importance to the present project are:
  - (a) The correlation coefficient between loadings on two factors.
  - (b) Cattell's coefficient of pattern similarity.
- 5. Possible effects of change in battery analyzed by test addition or deletion have been noted. Common factors may be added when tests are added or common factors may be deleted when tests are deleted. Addition of tests might help determine a simple structure while test deletion might leave the structure indeterminate.
- 4. Effects of selection of the group of subjects used in a study can change the extent of correlations between factors and the relative magnitude of loadings of the tests on the factors. In some extreme cases, factors may be added or deleted. The factorial composition of tests may change when groups of different levels of ability are being examined.

# B. Solutions to Similar or Related Problems

Two particularly relevent developments have been reported in the literature. In one, Mosier (17) has treated the case where loadings on a factor are assumed to be known and it is desired to locate an axis which has these loadings within small differences. A least-squares method of fitting actual loadings to the theoretical loadings were used. His solution resulted in the following matrix equation:

$$(A'A + \beta I)A = A'V$$

(II.

where A is the factor matrix on reference factors, A is a constant to be determined, A is the column vector with direction cosines of the desired axis, and V is a column vector with the theoretical loadings to be approximated. Since the labor of solution would be great, Mosier suggested as approximation obtained when A was set at zero.

$$\Lambda = (A'A)^{-1}A'V$$

(II.

This would result in direction numbers imstead of direction cosines, and it would be necessary to normalize the solution. Justification of approximation can be made in either of two ways: one, that  $\underline{\mathcal{L}}$  should be small; and two, that the restriction that the resulting vector be of unit length be discarded. This latter method of justification alters the problem to that of finding an axis for which the loadings are as nearly proportional to the theoretical loadings as possible. Equation II.2 is a solution to this problem.

Harold Hotelling (13) has dealt with the situation where there are not only several predictor variables to be combined in a regression equation but also several criteria which are to be combined so that optimum prediction can be obtained.

The matrix equations are:

$$Z = X_{\mathbf{C}} A_{\mathbf{C}};$$
 (II

$$(R_{i\beta}^{*}R_{ij}^{-1}R_{j\alpha} - \lambda R_{\alpha\beta})A_{\alpha} = 0 ; (II$$

$$A'_{\alpha} R_{ci\beta} A_{\beta} = 1; \qquad (II$$

where: X<sub>m</sub> is the matrix of standard scores on the criteria,

A<sub>m</sub> = A<sub>r</sub> is a column vector of optimum weights for the criteria,

Z is a column vector of optimally predictable criterion standard scores,

R<sub>1j</sub> is the matrix of predictor intercorrelations,

R<sub>2</sub> is the matrix of criterion intercorrelations,

R<sub>3</sub> = R<sub>3</sub> is the matrix of correlations of predictors with criteria,

\$\lambda\$ is an undetermined multiplier.

Once equations II.4 and II.5 are solved, the appropriate regression of the predictors for the criterion 2 can be obtained by usual methods where:

$$R_{iz} = R_{i,0}A_{\alpha}, \qquad (II.6)$$

 $R_{1Z}$  being a column vector of correlations of predictors with Z. Hotelling notes that equation II.4 may be simplified by treating the original criteria so as to obtain a derived uncorrelated set to be used in this solution. The matrix  $R_{O,B}$  is then the identity matrix and may be dropped. Equation II.4 is then in the form for solution for principal components.

Estelling's solution is of interest in the context of the present project in that he was matching optimally two separate sets of observations. We will be interested in matching factor loadings of tests rather than scores of individuals.

# III. General Conditions and Assumptions

# A. Assumption of Thurstone's General Theory of Factor Analysis

Since considerable controversy exists between several systems of .
factor analysis it was relt necessary to limit consideration of synthesis of factor analysis studies to within a particular factorial system, that developed by L. L. Thurstone (28). Of particular importance are the following basic assumptions made in Thurstone's theory of multiple factor analysis.

- Variability of scores on a test among members of a
  population can be accounted for by variability of
  underlying abilities among members of the population,
  plus errors of observation.
- 2. For any particular battery of tests, some of the abilities are common in that they contribute variability to scores on several tests and some abilities are unique to each test in the battery.
- A linear combination of ability scores is an adequate approximation to the actual mode of combination in producing test scores.

As a consequence of the choice of framework it would be expected that the method of synthesis would be more likely to be satisfactory for studies made within Thurstone's theory. Additional values might accrue if the method of synthesis were found to be usable for other factorial theories, but the reasonableness of the application would have to be separately evaluated.

# B. Effects of the Groups Tested.

One of the problems in synthesizing factorial studies is that of abstracting common meaning in spite of the fact that different groups of individuals have been used as the base for the several studies. Since these base groups are not usually selected as unbiased samples from the same population, it seems wise to consider that they may have different

parameters such as test and factor variances and intercorrelations. The base groups often are composed of the people most readily available for testing. Some restrictions may be imposed, but no real attempt is made to obtain an unbiased sample of a previously defined population. When several such catch-as-catch-can groups are involved, the statistics for these groups may well differ greatly. It is obvious that an adequate method of synthesis must be able to cope with these inter-group differences. One restriction that must be made in order to develop a method of synthesis is that the base groups do not differ so widely that the factorial pattern of tests change markedly. If performance on a particular test is more dependent on reasoning ability for grade school children but more dependent on perceptual speed for college students, the factor patterns would be different and use of results for this test would be misleading when attempting to synthesize the two studies. Wide differences between the groups must, consequently, be considered with skepticism.

# C. Defined Restrictions.

**日本社会の対象を表する。この会社会の対象の対象の対象にある。** 

In the light of the foregoing discussion and the material in Appendix A, it seems necessary that three areas of possible restrictions be considered and that necessary definitions be made.

1. Since the test variances are a function of the group of people who happen to take the tests, and the rows of the factor matrix are affected proportionately if the units of measurements of the tests are changed, it seems desirable to establish a common unit of measurement for each test taken by the several group. This requires a factor analysis of covariances rather than of correlations. When correlations are analyzed, a set of units are implied which yield unit variances for one particular group of people. When two groups are considered with different test variances, two different sets of test units of measurement are necessary to yield correlations for each of the groups. In Section III of the Appendix A, it is shown that a change in test unit of measurement results in proportionate changes in the factor loadings of the tests (See equation A.20).

As a consequence it will be defined throughout the present development that the factor matrices for the several groups will have been adjusted so that a common unit of measurement will have been used for each test across the several groups. The method for making the adjustment for units of measurement is outlined below. In this outline it is assumed that the factor analyses have been carried out on correlations.

a. Determine for each group a desired standard deviation on one test. These standard deviations are to be proportional to the raw score standard deviations on the test for the groups. The average of the standard deviations over the several groups is to be considered as a weighting factor for the test indicating the importance of that test in determining the synthesis of the factor analyses. If one of the tests is not to be relied on very heavily in synthesizing the studies, the average standard deviation of that test can be made low. Conversely, an important test can be given a high weight by using a large average standard deviation. In the illustrative pair of studies, for which the factor loadings were given in Table 1, all tests were assigned unit average standard deviations. Table 2 presents the raw score standard deviations in the two studies for each test. For the first test, the raw score standard deviations were 1.73 in Study A and 1.884 in Study B. The mean of these two standard deviations was 1.807. The adjusted standard deviations were obtained by dividing the raw score standard deviations by the mean:

> .9574 = 1.73/1.807; 1.0429 = 1.884/1.807.

In case this test were to have a weight different from unity, the adjusted standard deviations would have been multiplied by this weight. b. Multiply the factor loadings for the test in the reference factor matrix for each group by the desired standard deviation of the test for the group. Table 3 presents adjusted factor loadings for the example. Only the tests common to the two studies are included in this table. In obtaining these values, the factor loadings for each test were multiplied by the corresponding adjusted standard deviation. The loadings in Table 1 for Study A test 18 were multiplied by .9574, the corresponding adjusted standard deviation given in Table 2. Similarly, the loadings in Study B for this same test were multiplied by 1.0429. These steps can be stated in matrix algebra as:

$$F_{JEA} = D_{JA}F_{JMA}$$
, (III.1A  
 $F_{JMS} = D_{JB}F_{JMB}$ . (III.1B

The entries in the diagonal matrices are the adjusted standard deviations. The subscript  $\underline{J}$  is used to designate the tests with unit standard deviations in the separate studies and the subscript  $\underline{J}$  is used to designate the tests after adjustment of the standard deviations.

- c. Repeat steps a and b for each of the remaining tests.
- d. The resulting factor matrices with adjusted standard deviations will be used in all following steps.
- 2. In accordance with the results of section 4 and equation A.28 of Appendix A, the variances of scores on the factors will be permitted to vary between the two studies.
- 5. Since biased selection of the groups is anticipated as possible and the results of section h of appendix A indicate for this case that the correlations between the factors may differ for the several groups, no restriction of similar correlations between factors for the several groups will be imposed.

# D. Synthesis of Studies Two at a Time.

In order to simplify the problem to some manageable size, consideration will be given only to the case in which there are two studies to be synthesized. When there are more studies than two, it is hoped that complete synthesization can be accomplished by progressive synthesis of pairs of studies.

# E. Some Tests Common to Both Batteries, Other Tests in Either Battery.

It will be assumed that the batteries of tests in the two factor studies contain some overlap tests and some tests that appear in either of the batteries but not in both. A test will be considered as an overlap test only if 1) the test has not undergone editorial changes, 2) the time limits have not been changed, 3) instructions have not been changed, 4) test administration conditions are similar, and 5) scowing method is the same. Synthesis of the studies will depend on the overlap tests only.

Table 2
ADJUSTED TEST STANDARD DEVIATIONS

Test Code Numbers		Raw Score &	Standard Dev	Adjusted Standard Deviation			
Study A	Study B	Study A	Study B	Average	Study A	Study B	
18	1	1.73	1.884	1.807	.9574	1.0429	
9	26	8.04	9.487	8.764	.9174	1.0825	
5	23	. B.05	8,920	8.485	.9487	1.0513	
12	28	12.39	14.942	13,666	.9066	1.0934	
14	22	8.30	8.497	8.398	.9883	1.0118	
13	. 32	6.49	6.971	6.730	9643	1.0358	
14	34	5.91	6.529	6,220	.9502	1.0497	
15	33	7.22	8.695	7.958	.9073	1.0926	
ii	27	10.62	12.076	11.348	.9358	1.0642	
16+17	30	25.67	25.038	25.354	1.0125	.9875	

# LOADINGS ON REFERENCE FACTORS FOR OVERLAP TESTS

# WITH ADJUSTED STANDARD DEVIATIONS

# Matrix F<sub>JmA</sub>

tudy A	Study B	1	_11_	111	17	Y	<u> 71</u>
18	1	565	201	086	-010	067	-010
9	26	565 761	037	009	009	073	-009
5	23	541	-086	009	503	256	199
าร์	26	526	063	145	317	100	281
4	22	526 267	089	030	435	524	346
13	32	608	096	029	260	444	260
14	34	570	038	000	095	285	<b>676</b>
15	33	363	000	009	336	508	000
ŭ	27	730	047	384	309	Č09	000
16+17	30	528	633	104	. 122	011	017
		•	i				•

				<b>)</b>	htrix l	JICB		-							
Test Code	Test Code Rumbers Lordings on Reference Factors (Decimal points are omitted.)														
Study A	Study B	<u> I</u>	ĮŢ.	III	IV	<u> </u>	VI	VII	AIII.	IX	<u> </u>	XI	XII.		
18	1	439	-078	316	256	182	473	083	043	064	-010	050	121		
9	26	595	-024	071	249	167	401	-034	265	188	-065	297	-040		
. 5	23	191	C45	052	559	279	159	-œ8	193	470 488	097	212	O43		
12	28	253	-004	136	288	393	51:9	010	087	488	-087	510	055		
4	55	091	c3 <b>3</b>	050	622	. 231	800	014	137	384	-006	167	166		
13	32	331	054	063	609	193	238	017	186	379	034	215	016		
14	34	154	-047	Oh6	436	037	181	-058	616	179	023	283	-010		
15	33	රෙහිර	125	-02.5	836	-008	167	-046	161	251	-010	189	-016		
11	27	420	-059	114	311	0^5	512	011	164	315	092	496	094		
1617	<b>7</b> 0	777	007	MAS.	110	2.7	51.3	-000	CEL	17A	-ara	127	-001		

# IV. Congruence of Factor Matrices

# A. Definition of Congruence.

The term congruence is used in this development to indicate a lower level of precision of coincidence than is associated with its use in geometry. Bather than meaning that an exact fit of one matrix to the other has been obtained, an approximate fit is to be indicated by the term. Two matrices will be considered as congruent if they are generally similar, with only relatively small random differences.

In devising an index of extent of congruence, the concept of being able to replace the two matrices by an average matrix is convenient. Let the two matrices in which we are interested be the factor matrices  $F_{JrA}$  and  $F_{JrB}$ . Their average can be designated  $F_{Jrc}$ , and is obtained by the following equation in which lower case letters designate cell entries.

$$f_{Jrc} = (1/2) (f_{JrA} + f_{JrB}).$$
 (IV.1

In the index, the differences between each element of one matrix and the corresponding element in the average matrix is squared and these squared differences summed separately for each factor. This is done for both of the original matrices. Squares of the entries for each factor in the average matrix are then summed. Since it is important that each factor is congruent, a separate index of congruence, g,, is defined for each factor as the ratio of the sum of squares of differences for the factor to twice the sum of squares of average loadings for the factors:

$$B_{T} = \frac{\sum_{I} (f_{JrA} - f_{Jrc})^{2} + \sum_{I} (f_{JrB} - f_{Jrc})^{2}}{2 \sum_{I} f_{Jrc}^{2}}.$$
 (IV.2)

The denominator is doubled so as to balance the number of values summed in the numerator and the denominator. This index might best be described as a measure of the extent to which it would be sensible to replace the factor from the two studies by the average. A low value would indicate that this would be possible.

When equation IV.1 is substituted into equation IV.2, it is found that

$$\mathcal{E}_{r} = \frac{\frac{\Gamma}{J} \left( f_{JrA} - f_{JrB} \right)^{2}}{\frac{\Gamma}{\Gamma} \left( f_{JrA} + f_{JrB} \right)^{2}}.$$
 (IV.3)

The index of congruence for each factor r is then the ratio of the sum of squares of the differences between loadings for the tests on the factor in the two studies and the sum of squares of the sum of loadings for the tests on the factor in the two studies.

It is to be noted that the closest agreement between factors in the two studies, and thus maximum congruence, occurs when the index of congruence,  $g_{\chi}$ , is a minimum. A coefficient of congruence,  $g_{\chi}$ , for each factor is developed in Appendix B:

$$g_{r} = \frac{\sum_{J} f_{JrA} f_{JrB}}{\sqrt{(\frac{1}{2} f_{JrA}^{2}) (\frac{1}{2} f_{JrB}^{2})}}$$
(B.19)

The relation between g and g was found to be

$$g_{r} = \frac{1 - \phi_{r}}{1 + \phi_{r}}$$
 (B.18)

The coefficient of congruence,  $\phi_{\mathbf{r}}$ , has similar properties to a coefficient of correlation approaching a maximum of unity for the most precise congruence and a lower limit of zero for the least precise congruence.

Table 4 gives the factor loadings of the overlap tests in the two studies in the example for the six factors corresponding to the six maxima of the coefficient of congruence. Inspection of the loadings on factor B reveals only small differences between the two studies. The coefficient of congruence, \$\overline{\psi}\_r\$, for factor B is .999984 while the index for congruence, \$\overline{\psi}\_r\$, for factor B is .999984 while the index for congruence, \$\overline{\psi}\_r\$, is .000098. High congruence also occurs for factors A, D, and E. The congruence for factor C is moderately high but distinctly below that for those noted above. The congruence for factor F is definitely low so that this factor will not be considered as a congruent factor. The congruent opace between the two studies is therefore defined by the five factors A-E.

Table 5 gives the transformation matrices  $T_{mrA}$  and  $T_{MrB}$  used in rotation from the reference axes of  $F_{JmA}$  and  $F_{JMB}$  of Table 3 to the congruent factors of Table 4 by equations B.1A and B.1B. The columns of these transformations have been determined so as to maximize the coefficients of congruence (or, it may be stated, to minimize the indices of congruence). Once the reference frame of congruent factors has been determined, joint

rotations within the congruent space are possible. These rotated factors will also have high coefficients of congruence. Appendix B gives the mathematical development of the solution and Appendix C gives the computational procedures.

# B. The Non-Congruent Spaces

In the example there is a five dimensional congruent space. Study A has six factors which leaves one dimension not included in the congruent space. Table 6 gives the direction cosines for the dimension in this study which is uncorrelated with the congruent space. This dimension constitutes the Study A non-congruent space. Study B has a seven dimensional non-congruent space. The axes given in this space are mutually orthogonal and, as a set, orthogonal to the congruent space. In general, it would be expected that only a portion of the space in each of a pair of studies would be congruent, leaving a remainder of the space in each study non-congruent with the other study. This situation can arise from each study involving as a common factor some mental function not included in the common factor space of the other study. A second possibility is that the same factor might be involved in both studies, but the overlap tests in the two studies would not include tests adequately leaded with this factor to establish the congruence.

# C. Meaning of un Observed Congruence

A basic assumption underlies the general attempt to coalesce results from several factor analysis studies. This assumption is: If a mental function is represented by a factor in each of several studies, the factor loadings of the tests should be the same in these studies. The invariance of factor loadings, then, becomes a necessary condition for identity of factors in two studies. It has been found, however, to be both reasonable and necessary to qualify this condition when there is likelihood of between group changes in the mental functions underlying differences in test performances within the group: of people on which the factorial studies are based. Even so, the invariance of factor loadings remains as a basic principle. Congruent factors between two studies satisfy the necessary condition of invariance of factor loadings. These two factors may represent the same mental function.

Before concluding, however, that congruent factors always represent the same mental function in two studies, it is important to ask whather factorial invariance is a sufficient condition. The answer is no. invariance of factor loadings is not a sufficient condition to identity of mental functions. Consider, for example, a test composed of verbally stated computation problems. This test, in a comprehensive study, might have loadings on a verbal factor and a computation factor. Suppose that this test is common to two smaller studies and is the only overlap test in the verbal and computation domains. Suppose, further, that in Study A there are several other verbal tests but no other computation tests, and that in Study B the reverse is true, there are other computation tests but no verbal tests. In Study A the common factor space will include the verbal factor but not the computation factor. The reverse will be true for Study B for which the common factor space will include the computation factor but not the verbal factor. Remember that our verbally stated computing test had loadings on both of these factors but was the only overlap test in these domains between the two studies. In this situation it would appear that the verbal factor of Study A was congruent with the computation factor of Study B. The mental functions, however, would not be identical. While this example may seem extreme, it demonstrates the proposition that invariance of factor loadings for a limited set of tests is not a sufficient condition for identity of factors between two studies. A number of other situations may also lead to congruence of factors without identity of mental functions.

The congruence of factors is, then, a necessary but not a sufficient condition for identity of factors between two studies. The credence, as to identity of factors, one can place on an observed congruence of factors between two studies depends on the extent of data on which the congruence is based. If the overlap tests in two studies are few in number and of limited variety for each mental function, very little confidence as to identity of factors can be placed on an observed congruence. In the example there are only ten overlap tests while there are six common factors and twelve common factors in the two studies. If the ten tests covered a ten dimensional sub-space out of the twelve dimensions of Study B, perfect congruence would be found for all six factors in Study A. Mathematics

indicates this would be necessarily true. Suppose, even, that the ten tests were not identical between the two studies but were artificially matched. Perfect identity of factor loadings still could have been obtained. The congruent factors would be artificial.

In order to avoid this situation, in so far as possible in the method for obtaining congruent factors, the principal axes for the overlap tests are found in each study separately and minor axes are eliminated until the dimensionalities are reduced somewhat below the number of overlap tests. The extent that this is possible depends on the set-up of the two studies and choice of overlap tests. In the illustration, the dimensionalities used were six and eight principal axes in the two studies. This represented are reduction for Study A and a moderate one for Study B. Artificial congruence will still be a major problem in interpretation of the results.

AND THE PARTY OF T

THE CONTRACTOR OF THE PROPERTY OF THE PROPERTY

Several rotations of axes within the congruent space were tried in order to find what dimensions might be more strongly determined than other dimensions. Table 7 gives the resulting factor loadings. Table 8 gives the corresponding transformations from the reference axes. Loadings between .15 and .24 have been singly underlined, loadings of .25 or more have been doubly underlined. Factors d and e have only one test each with underlines. In consequence very little confidence can be placed in the meaning of the congruence of these factors between the two studies. Factors a and b each have four tests with underlines and much more confidence can be placed in their congruence. Factor o occupies an intermediate position. As a result of our judgments of confidence in an observed congruence, some of the observed congruent factors in which we have low confidence can be reclassified to the non-congruent space. Factors d and e might be so transferred. The remaining factors, then, constitute the space on which confidence of the meaning of the congruence can be placed. Table 9 presents the factor loadings of the non-overlap tests in the two studies on factors a, b, and c. These loadings and those of the overlap tests may be inspected for interpretation of the two studies combined.

Whenever strong indications of identity of factors is desired by the use of overlap tests, we may conclude, it is necessary to include in the study plans provision for an adequate number of overlap tests. These tests should be varied as to form while depending on the same factors. Several parallel forms of each of a few tests will not be an improvement over

including only one form of each test. These parallel forms will not provide distinctive data on which to base confidence in an observed congruence. Busically, then, confidence in observed similarities between results of two studies depends on adequate experimental design. Methods of analysis can then be of assistance.

Matrices For and Form

					-	24	-							
	Factor F	Study	17	8	01	8	.05	.8	6	ន	さい	8	.459717	370129
	Pact	Study	3	6	16	કુ	.10	8	કુ	03	8	<b>す</b>	24.	Ŋ.
rs Factor E	ار ا	Study	8.	8	20.	ន	01	8:	05	15	21	11.	019665	.000165
	·	Study	l '	8	ဝ့်	ਰੋ.	01	8	05	15	8	-17	83.	§.
t factor	O Z	Study B	ð	•.05	#.	8	8	8	8	17	สฺ	.01	999875	290000
234	Page 1	Study	ਰੋ.	05	Ħ,	88.	ş	કં	න	17	ૹ઼	10.	86.	8
	or C	Study Study A B	03	13	ਰੋਂ.	† 0	50	01	03	#.	8	.16	939811	.031028
	Fact	Study	₹O	÷1	දු දි	9.	<b>ਰ</b>	8	§.	.13	<u>.07</u>	.17	.93	.03
	or B	Study B	41.	5	ଅ'	8:	φ <u>ε</u> ••	-,23	8,	26	-26	EL.	186686	800000
\$		Study A	4.	દ	63.	81	٥: ا	23	81	સું '	ઌૣૺ૽	¥.	ķ.	ġ.
•	T TOOL	Prace Pra Prace Prace Prace Pra Prace Pra Pra Prace Pra Pra Pra Prace Pra Pra Pra Pra Pra Pra Pra Pra Pra Pra	8.	:: ::	<b>₹</b>	à i	10.	si,	8	-17	51.	.43	979633	000058
ļ	200	Study	O.	7 7	4 6	3 8	70.	7,	Q.	21:	j.	3.	χ <b>.</b>	, ox
24 27 Pa	0 10	Stricty	r-t\	0 KG	n a	3 6	\ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \	i K	\$ 1	3 6	, v	. 00		
Teat Toly N. Wales	1	A Congra	స్తే :	بر در ب	٠, ڏ <u>.</u>	ų	÷ ,	7 4	† (	J:	717	TOUR !	* 13.4	.} 'to'

 $^*\!\!f_T$  and  $g_T$  were computed from factor leadings carried to four decimal pluces.

-24-

Table 5
TRANSFORMATION TO CONGRESSIT FACTORS

### STUDY A

# Matrix T\_\_\_

Reference	Congruent Factors												
Factors	A	_B	C	D		<u> </u>							
1	.4786	.1156	2175	0188	.0372	±.1249							
11	.3765	.0829	<b>3400</b>	0972	.3290	.1473							
111	5678	.7978	.1197	.4222	8345	.6199							
<b>1Y</b> .	0015	4349	.5502	.2500	.2159	5081							
· 🔻	.0148	3283	.0456	4994	4410	3549							
<b>vi</b>	4190	2634	5835	.7162	.3301	.3955							

# STUDY B

# Matrix T<sub>MrB</sub>

Reference		Co	gruent :	<b>Pactors</b>		
<u> Factors</u>	<u>A</u>	B	C	D	E	<u> </u>
Ĭ	.1584	0729	6344	0656	0529	3057
II	.1545	1961	.1053	1769	.1405	1439
ΙΙΥ	1594	.1217	2053	.0816	1206	.6394
IV	.1097	3825	.1095	3203	1741	.1678
Ą	.3365	3786	.0097	.0739	.6545	.1293
VI.	.6015	•5972	.6999	1570	.0429	0269
VII	1792	.0149	1208	.1086	1020	.1910
YIII	.3478	1368	2637	1773	.1952	0268
IX	1983	2664	.0524	•5074	•0942	2988
X	1803	.1450	.3323	.4671	1605	5283
XI	3897	.4004	0623	.3630	4435	.1459
XII	4391	.0301	1371	.4019	2581	.2066

<sup>\*</sup>Congruent Factor F had such a low coefficient of congruence that it was eliminated from the congruent space.

Table 6
DIRECTION COSINES OF NON-CONGRUENT FACTORS
STUDY A

# Matrix A muA

Reference Factors	Non-Congruent Factor
<b>i</b>	0827
11	•759i
111	,1610
14	<b>~.</b> 1653
₩	.4628
vi	,3865

# STUDY B

# Matrix A MUB

Reference	•	Non-Congruent Factors										
Factors	1_	5	3	4	5	6	7					
Ι	0043	0592	0244	0399	1537	2042	4464					
II	.1899	2934	.8158	2423	1728	.0619	1633					
III	5234	0476	.2507	3179	.6857	Ö	0					
17	0289	1110	1946	.0522	.3070	.1871	0515					
Ą	.3235	4013	2307	4610.	.2892	.2273	1741					
ΥI	.1315	•1639°	.0913	.0437	.2250	.0289	1650					
AII	0064	~.ceh4	•3565	.8911	.2215	0776	0782					
VIII	.0361	.0407	.1576	.0511	.0332	.1205	.8075					
ZZ	2141	.5799	.1935	0095	0545	.2317	1008					
X	2465	4920	~.0268	0168	1399	4656	.2105					
XI	.0662	5271	0?33	.0006	1280	.7266	0					
XII	.7552	.0981	.0179	1858	.3907	5531+	.0424					

Table 7

# LOADINGS OF OVERLAP IESTS OR ROTATED CONCRUENT PACTORS

Matrices Fir and Firm

actors	Factor e	Study B	នុះ	ខ្ម	8	8	8.	 20:	8	8	<u>ب</u>
		Study A Study	ន់ខ	કં ક	ð	ક્	9.	05	8	8	સ
	Pictor d	Study B	80.7	ន់	<b>ત</b>	01	8	3	8	.35	8
		Study A	8.	9.	.13	70.	01	3.	8	35	8
Congruent Factors	٠ ١	Study B	9.8	, e	_ ਸ	119	듸	7.	4	8	8
Rotated (	rector	Study A Study B	<b>4</b> 8	; <b>≓</b>	13	S	.13	r.	3.	8	8
Loadings on Rotated	Factor b	Study B	01.	ક	e:	.31	319	03	8	jo.	8
		Study A Study B	01.	કે કો	<u>ن</u>	87	÷.		ន់	-07	8
	rector a	Study B	21	<b>1</b> 6	ಕ	.03	ij.	:23	8	<b>d</b> .	ક
		Study A Study B	51.	įį	%	03	o.	53	કં	10	8.
Test	Code Numbers	tudy A Study B	чģ	3 23	සූ	ଖ		तं	33	2,1	ន
		Study A	ရု ၀	, iv	ឌ	. <del></del> †	<b>1</b> 27	† <del>*</del>	i.	#	16+17

# Table 8

# TRANSFORMATION TO ROTATED CONGRUENT FACTORS FROM REFERENCE FACTORS

# STUDY A Matrix T<sub>mzA</sub>

Reference	Rotated Congruent Factors						
Factors		þ	¢	. а	•		
1	.4368	0732	0863	0342	0173		
11	2085	0435	.0441	1374	.4972		
111	3343	1220	0731	.9301	\\369		
iv	6100	.5790	.2769	.0906	.5285		
*	.1048	-,5297	.7422	0567	3316		
<b>ty</b>	.2558	.7629	-,8453	0184	1750		

# STUDY B Matrix T<sub>MzB</sub>

Reference	;	Rotated	Congruent	Factor	
Factors	8.	ъ	c	đ	
I	.7558	1017	2149	1783	4411
ıī	0272	0270	.1768	1925	.1504
III	.1374	0370	1681	.0799	2273
14	.0210	0886	.558 <b>5</b>	1060	0779
Ÿ	.0466	3592	2186	4190	.4706
ΥÌ	4292	3465	.2638	.2205	.6306
VII	.0395	.0608	0935	.0880	1629
VIII	1,256	0925	0560	-,3056	0216
IX	2188	.6453	1669	.2038	.1270
X	4600	.3954	0200	.4952	.1780
XI	1090	.0604	1797	-5336	3135
XII	0799	.3062	1889	.3479	2725

Table 9

# LOADINGS OF NON-OVERLAP TESTS ON ROTATED CONORDENT FACTORS

STUDY A

RITERY &

MATRIX P<sub>JRA</sub>

MATRIX P<sub>JEB</sub>

	JAA	Jan				
Test Code Numbers	Loadings on Rotated Congruent Factors a b c	Test Code Numbers	Loadings on Rotated Congruent Pactors a b c			
1	.2836 ,2093 -,1605	2	.111513531424			
3	03370165 .0585	3	.34720339 .0275			
10	.0945 .03161694	Ĭ.	.3123 .08090355			
2	.084710560417	2 3 5 6 7 8 9	13020865 .3519			
2 6	.0516 .06721735	é	.210011411112			
7	.06120620 .0067	7	0406 .04620895			
7 8	0730 .00460341	ė	.3794 .01241869			
		ğ	.473803780925			
		10	.0269 .0315 .0026			
	*	n'	0461 .10360092			
		12	06760875 .0650			
		13	.4293 .06570264			
		14	,3887 ,0361 -,1263			
		15	,2163 .0807 .1397			
		16	.0024 .2867 .0842			
		17	.0825 .0062 .3041			
		ī8	0418 .2174 .2423			
		19	.0362 .27911138			
		ēό	.0406 .06530291			
	•	21	.0349 .23091105			
	<b>i</b> '	24	.1714 .09750994			
	<b>₽</b>	25	.4152 .0729 .0077			
		29	.0179 .1987 .1495			
		31	.05280166 .0337			
		55	.26510655 .0111			
	`	<b>3</b> 5 36	7000. \$\$40. \$850.			
	•	37	07160548 .3249			
		38 38	1311 .1797 .1005			
		<del>3</del> 9	.05820939 .2909			
		40	2435 1254 - 1043			
		. 41	.110706171515			
	•	. <u>115</u>	035105632138			
		43	.1468 .0328 .0594			
		44 44	0799 .03782053			
		44	1012A - 1001e - 15029			

# References

- Burt, Cyril. Factors of the Mind. London: University of London Press, 1940.
- 2. Cattell, R. B. "Parallel proportional profiles" and other principles for determining the choice of factors by rotation. <u>Psychometrika</u>, 1944, 9, 267-283.
- Cattell, Raymond B. The description and measurement of personality. London: Harrap and Company, 1947.
- 4. Cattell, Raymond B. Confirmation and clarification of primary personality factors. Psychometrika, 1947, 12, 197-220.
- 5. Cattell, Raymond B. The primary personality factors in women compared with those in man. Brit. J. Psychol. Statist. Sect., 1948, 1, 114-131.
- 6. Cattell, Raymond B. A note on factor invariance and the identification of factors. Brit. J. Psychol. Statist. Sect., 1949, 2, 134-139.
- Cattell, Raymond B. r and other coefficients of pattern similarity. <u>Psychometrika</u>, 1949, <u>14</u>, 279-298.
- 8. Fiske, p. W. Consistency of factorial structure of personality from different sources. Amer. Psychol., 1948, 3, 350.
- 9. Greenall, P. D. Two criticisms. Brit. J. Psychol. Statist. Sect., 1943, 1, 64.
- 10. Holzinger, Kurl J. and Harran, Harry H. Relations between factors obtained from certain analyses. J. educ. Psychol., 1957, 28, 321-345.
- 11. Hotelling, Barold. Analysis of a complex of statistical variables into principal components. <u>J. educ. Psychol.</u>, 983, 24, 417-441 and 498-520.
- 12. Hotelling, Herold. Simplified calculation of principal components.

  Psychopetrika, 1935, 1, 27-35.
- 13. Notelling, Marold. The most predictable criterion. J. educ. Psychol., 1935, 26, 139-142.
- 14. Kelley, T. L. Essential traits of mental life. Cambridge, Mess.: Harvard Univ. Press, 1935.
- 15. Ludin, A. A note on "criterion analysis". Psychol. Rev., 1950, 57, 54-57.

# References, continued

- 16. Meyer, Lorenz A. The invariance of factorial composition of a test. PhD dissertation on file with The University of Chicago Libraries. 1943.
- 17. Mosier, Charles I. Determining a simple structure when loadings for certain tests are known. Psychometrika, 1939, 4, 149-162.
- 18. Reyburn, H. A. and Taylor, J. G. On the interpretation of common factors.

  Psychometrika, 1943 8, 53-64.
- 19. Reyburn, H. A., and Raath, M. J. Simple structure: A critical examination.
  Brit. J. Psychol. Statist. Sect., 1949, 2, 125-133.
- 20. Saunders, D. R. Factor Analysis: I, Some effect of chance error.

  Psychometrika, 1948, 13, 251-257.
- 21. Smart, R. G. The variation in pattern of factor loadings.

  J. educ. Paychol., 1937, 28, 55-64.
- 22. Stephenson, W. The inverted factor technique. Brit. J. Psychol., 1936, 26, 344-361.
- 23. Thorson, G. H. The factorial analysis of human ability. New York:
  Houghton Mirriin Company, 1946.
- 24. Thurstone, L. L. Primary Mental Abilities. Psychometric Monograph Number 1. Chicago: University of Chicago Press, 1938.
- 25. Thurstone, L. L. The perceptual factor. Psychometrika, 1938, 3, 1-17.
- Thurstone, L. L. Experimental study of simple structure. <u>Psychometrika</u>, 1940, 5, 153-168.
- 27. Thurstone, L. L. and Thurstone, Theira Gwinn. Factorial studies of intelligence. Psychomet: 'c Monograph Number 2. Chicago: University of Chicago Press, 1941.
- . 28. Thurstone, L. L. <u>Multiple Factor Analysis</u>, Chicago: University of Chicago Press, 1946.
- 29. Young, Gale, and Householder, A. A. Factorial invariance and significance.

  Psychometrika, 1940, 5, 47-56.

#### APPENDIX A

#### THEORY OF MILITIPLE FACTOR ANALYSIS OF COVARIANCE

Multiple factor analysis has been developed using a restrictive definition concerning the unit of measurement for each variable, be it a test or a factor, such that unit variance is obtained for the group of people on whom the study is based. The initial equation relating test scores to factor loadings and factor scores is commonly written in terms of standard scores for tests and factors. The units of measurement are treated as "floating", taking such values as yield unit test or factor variances for the group concerned. It is the purpose of these notes to derive the theory of multiple factor analysis when the variances are permitted to have any values, thus permitting the units of measurement to be otherwise defined. This is a necessary step when several groups of people are used for different factor analyses of the same tests, since these different groups are likely to have different variances of scores when the units of measurement of the variables are held constant for the several groups.

#### 1. General Factorial Equation for Covariances.

Equation A.1 is the usual initial linear equation used in factor analysis. In it and subsequent equations, capital letters are used to designate matrices, and lower case letters are used to designate call entries in the corresponding matrices. Subscripts are attached both to matrix designations and to call entry designations for convenience in keeping track of the variables involved. Particular notation used in this section is as follows:

N - number of people in the group,

X = watrix of scores,

C = matrix of variances and covariances,

F - metrix of factor leadings,

i - subscript designating individual person,

J and k - an elternate pair of subscripts designating tests, and

p and q = an alternate pair of subscripts designating factors.

Equation A.1 is then:

These scores are deviation scores, but no restriction is placed on their variances. The restriction of deviation scores is of nr importance since corrections for means on the factor scores leads to correction for means of the test scores without requiring a change in the factor matrix. As a consequence we shall ignore any difference in means between groups, assuming that within-group deviation scores are used in all cases. (This action would not be warranted if groups were to be combined.)

Equation A.2 and A.3 are the usual ones for tables of variances and covariances:

$$c_{jk} = (1/n)x_{ji}x_{ki}^{r}$$
; (A.2)

$$c_{pq} = (1/n)x_{pi}x_{qi}'$$
 (A.3)

When equation A.1 is substituted into equation A.2,

$$C_{jk} = (1/\pi)F_{jp}X_{pi}X_{qi}^{i}F_{kq}^{i}$$
(A.4)

Eliminating the factor-score matrices of equation A.4 by noting that the right-hand side of equation A.3 is involved in equation A.4,

$$C_{jk} = F_{jp}C_{pq}F_{kq}^{\dagger} \tag{A.5}$$

Equation A.5 is the general factorial equation for covariances.

#### 2. Transformation of Factors.

In this section an inductive process in which it is assumed that transformations are possible seems appropriate, and the proof consists of a demonstration that this assumption yields consistent results.

Let T<sub>pr</sub> be any square matrix of an order equal to the number of factors and for which an inverse exists. It is our assumption that this matrix can transform one factorial matrix, with factors p (elternate subscript g), into another factorial matrix, with factors r (elternate of subscript g), by equation A.60.

$$\mathbf{F}_{\mathbf{jr}} = \mathbf{F}_{\mathbf{jp}} \mathbf{T}_{\mathbf{pr}} . \tag{A.6a}$$

(A.61

Eubstitution of equation A.6b into equation A.1 yields:

$$X_{j1} = Y_{j1}Y_{j1}^{-1}X_{j1} \tag{A.7}$$

Let the following definition of transformation of scores be made:

$$x_{ri} = x_{pr}^{-1}x_{pi} . (A.8a)$$

$$X_{pi} = T_{pr}X_{ri}. (A.8b)$$

Equation A.7 then becomes:

$$\mathbf{x}_{,ii} = \mathbf{F}_{,ir}\mathbf{x}_{,ri} . \tag{A.9}$$

Equation A.9 reproduces equation A.1 with factors <u>r</u> replacing factors p. Equation A.5 can be immediatedly re-written for factors <u>r</u>:

$$C_{ik} = F_{ir}C_{rs}F_{ks}, \qquad (A.10)$$

Similarity of equation A.8b to equation A.9, if matrix  $T_{pr}$  is considered to be a factorial matrix for factors p in terms of factors r, permits equation A.10 to be re-written.

$$C_{pq} = T_{pr}C_{rg}T_{qe}^{\dagger}, \qquad (A.11b)$$

$$C_{rs} = T_{pr}^{-1} C_{pq}^{T_s^{s-1}}$$
, (A.11a)

(The first of these equations is imbelled A.11b and the second A.11a so as to be consistent with equations A.6 and A.8. The <u>a</u> equations relate factors <u>r</u> to factors <u>p</u> and the <u>b</u> equations relate factors <u>p</u> to factors <u>r</u>.) Substitution of equation A.11a into equation A.10 yields:

$$c_{jk} = F_{jr}^{-1} c_{pr}^{-1} c_{qs}^{r_{s}^{-1}} ks$$
 (A.12)

Noting that the first two matrices on the right of equation A.12 reproduce the right-hand side of equation A.6b and that the last two matrices

of equation A.12 are the transpose of the right-hand side of equation A.6b, the corresponding substitutions yield equation A.5. Thus the system is internally consistent and the transformation of factors is possible.

Since there are an infinitely large number of matrices which satisfy the restrictions that they are of the order equal to the number of factors and possess an inverse, there are as many possible sets of factors which satisfy equation A.5, or A.10. This is the same problem as encountered in the normal factor analysis of correlations, and the solution proposed by Thurstone, that transformation be to a simple structure, is appropriate for the factor analysis of covariances. Actually, factor analysis of correlation is a special case of factor analysis of covariances in which the additional definition is imposed that the variances (diagonal entries in the C matrices) be unity. Existence of a simple structure, however, will not completely solve the problem for factor analysis of covariances. There remains a problem of changes in the size of the units of measurement for the factors.

Let there be a change in size of unit of measurement from factors  $\underline{r}$  to factors  $\underline{R}$ . This results in a proportional change of all scores on each factor. Equation A.13 accomplishes these proportional changes where the diagonal entries in  $\underline{D}_{\underline{R}}$  are the constants of proportionally.

$$X_{ri} = D_{ri}X_{Ri} . (A.13)$$

In order to simplify the algebra, the scores on factors  $\underline{r}$  were considered as proportional to the scores on factors  $\underline{R}$ . Equation A.13 is similar to equation A.8b, in that in A.13 factors  $\underline{r}$  are being transformed to factors  $\underline{R}$  by the matrix  $D_{\underline{r}R}$  just as factors  $\underline{p}$  were transformed to factors  $\underline{r}$  by the matrix  $\underline{T}_{\underline{p}\underline{r}}$  in equation A.8. Equation A.6a is then re-written:

$$\mathbf{F}_{\mathbf{JR}} = \mathbf{F}_{\mathbf{Jr}}^{\mathbf{D}}_{\mathbf{rR}} . \tag{A.14}$$

In this case, then, the columns of factor leadings on factors R are proportional to the leadings on factors r. The configuration of zero factor leadings is unchanged by this transformation and the simple-structure remains. Equation A.lla may be re-written.

$$C_{RS} = D_{rR}^{-1} C_{rs}^{-1} D_{sS}^{-1}$$

(1.15

Since the factors are not directly observed, no experimentally determined set of units of measurement exists. This poses a dilemna, for any set can be used and the simple structure will remain. For any single factor analysis the factor units of measurement can be left as unknown and a restriction placed on the variances of the factor scores. Thus the diagonal entires in the covariance matrix C<sub>RS</sub> can be defined as some constant such as unity. When several factor analyses on different groups are being considered, this simple solution is inappropriate, for the resulting units of measurement may be of different sizes for the several groups.

#### 3. Transformation of Tests

Two types of transformations of the tests are of interest. In theory, the type in which weighted sums of the tests are taken as new variables is the more general and includes as a special case the second type in which the units of measurement of the tests are changed so as to change the scores proportionally. Let it be desired to obtain variables h from tests j by a weighting matrix  $W_{h,j}$  in accordance with the equation:

$$X_{hi} = V_{hi}X_{hi}$$
 (A.16)

If equation A.1 is pre-multiplied by Whi:

$$W_{i,1}X_{1i} = W_{i,1}X_{0i}. (A.17)$$

The right-hand side of this equation can be simplified by the following definition:

$$F_{hv} = W_{h,i}F_{,ip}. \tag{A.18}$$

When equations A.16 and A.18 are substituted into equation A17:

$$X_{hi} = F_{hp}X_{pi}$$
 (A.19)

Equation A.19 is similar to equation A.1 and all of the derivations of sections 1 and 2 apply to the variables h. It is to be noted that there was no change in the factors or factor scores.

If a change is to be made in the units of measurement of the tests so that test i becomes test J, the weight matrix of equation A.16 becomes a diagonal matrix so that the scores on each test are changed proportionally. The weight matrix can be designated by  $D_{J,j}$  for this case. Then, if the subscript J is substituted for h and the matrix  $D_{J,j}$  is substituted for  $W_{h,j}$ , equations A.16-A.19 give the relations for the tests with new units of measurement. A result of this is that the analysis may be carried through with one set of units for the tests and the factor matrix can be transformed for a new set of units. For sake of later convenience equation A.18 is re-written with the necessary changes indicated above:

$$\mathbf{F}_{\mathbf{J}\mathbf{p}} = \mathbf{D}_{\mathbf{J}\mathbf{A}}\mathbf{F}_{\mathbf{J}\mathbf{p}}.\tag{A.20}$$

#### 4. Effects of Between-Group Differences

In this section it is assumed that the same battery of tests and set of factors are involved for two or more groups. Equation A.1 is assumed to hold for each group individually. It is relatively obvious that within-group means accres can be ignored provided that the groups are not to be combined. Deviation scores within each group can be obtained without changing the factorial matrix, and therefore will be used. Let there be two groups, A and B. Equation A.1 then expends to:

$$x_{jiB} = F_{jr}x_{riB}.$$
 (A.21B)

(Capitel letters are used at the end of equation numbers to indicate the groups, A or B, to which the equation applies.) It is assured that the groups are sufficiently similar that the factorial nature of the tests remain unaltered. If this is true, then for any particular individual it should not matter within which group he is considered so far as the factorial equation for his scores is concerned. The factorial equations

should therefore be parallel, with the same factor matrix F ... as shown in equation A.21. The matrices X jiA, X jiB, X riA, and X riB differ only with respect to which people are included in the groups A and B. By equation A.5:

$$C_{jkA} = F_{jr}C_{rsA}F_{ks},$$

$$C_{jkB} = F_{jr}C_{rsB}F_{ks}.$$
(A.22A)

An implicit assumption in equation A.21 and A.22 is that a single unit of measurement exists for each factor and is common for the two groups. As previously noted, the unit of measurement cannot be observed for factors. Consequently it is necessary to derive the relations between factor matrices for the two groups when the group factor variances are defined to be some constant. Employing the transformations of equations A.13-A.15 and noting that the matrices D<sub>rR</sub> are particular to the groups:

$$X_{RiA} = D_{rRA}^{-1} X_{riA};$$

$$X_{RiB} = D_{rRB}^{-1} X_{riB};$$

$$(A.23B)$$

$$F_{JRA} = F_{Jr}^{D} D_{rRA};$$

$$(A.24A)$$

$$F_{JNB} = F_{Jr}D_{rRB}$$
; (A.248)

(A.24B

$$C_{ASA} = D_{TRA}^{-1} C_{TSA} D_{SSA}^{-1}; \qquad (A.25A)$$

$$C_{RSD} = D_{rRB}^{-1} C_{rBB}^{-1} C_{rSB}^{-1} . \qquad (A.25B)$$

It is to be noted in equations A.24 that the factor matrices for factors R now have a subscript designating group. When equations A.24 are solved simultaneously so as to eliminate Fir

$$\mathbf{F}_{jRB} = \mathbf{F}_{jRA} \mathbf{D}_{rRA}^{-1} \mathbf{D}_{rRB}^{-1}. \tag{A.26}$$

It is possible to combine the two diagonal matrices by defining:

$$D_{RAB} = D_{rRA}^{-1}D_{rRB}^{-1}$$

(A.27

Then

(A.28

Thus the two factor matrices are proportional by columns. It is of. interest that the diagonal entries in  $\underline{D}_{RAB}$  are the ratios of the wariances on factors  $\underline{r}$  for group  $\underline{B}$  to those for group  $\underline{A}$ .

Even though the covariance matrices  $C_{RSA}$  and  $C_{RSB}$  have the same diagonal values, the off-diagonal entries will differ due to sampling effects, either from random sampling or selective sampling. If, as might be usual, the factor variances were to be defined as unity, these covariance matrices would become the correlation matrices between the factors for the two different groups. Under the assumption that the factors have some reality (which is necessary for any of this development and any hope that the same factors are operative for the two groups), it would be expected that the correlation matrices for the two groups would differ.

#### APPENDIX B

#### DETAILS OF MINIMUM SOLUTION FOR INDEX OF CONGRUENCE

In equation IV.5 the formula for the index of congruence for each factor was given as:

$$g_{r} = \frac{T_{JrA} - f_{JrB})^{R}}{\frac{T_{JrA} + f_{JrB}}{T_{JrB}}}.$$
 (IV.3)

It is desired to determine the factors so as to minimize  $\mathbf{g}_{\mathbf{r}}$ .

Consider that the analyses for the two studies have been factored to uncorrelated reference factors for the groups involved. Any set of orthogonal factors may be used whether they are the original factors obtained or an orthogonal rotation from the original factors. Let m represent the reference factors for study A and M represent the reference factors for study B. Then by equation A.6a:

$$F_{JrA} = F_{JmA}^{T}_{mrA}$$
 (B.1A)

$$\mathbf{F}_{\text{TrR}} = \mathbf{F}_{\text{TMR}} \mathbf{T}_{\text{MrR}}. \tag{B.1B}$$

(The last letter of equation number indicates to which study the equation applies. When both studies are involved, no letter will be used.) Writing these equations in summational notation:

$$f_{JrA} = \sum_{m} f_{JmA} t_{mrA}, \qquad (B.2A)$$

When equation B.2 are substituted into the index of congruence in equation IV.3,

$$g_{r} = \frac{\sum_{J} (\sum_{m} f_{JmA} t_{mrA} - \sum_{M} f_{JMB} t_{MrB})^{2}}{\sum_{J} (\sum_{m} f_{JmA} t_{mrA} + \sum_{M} f_{JMB} t_{MrB})^{2}}.$$
 (B.3)

Inspection of equation B.5 reveals that all t's can be multiplied by some constant without altering the value of g. This is true because the constant enters into each term in a similar manner, can be factored out, and then cancelled from numerator and denominator. Therefore, it is possible to derine a condition that the denominator equals some constant K without limiting the generality of the solution. The equation giving this condition is

$$\theta_{m} = \frac{\Sigma}{J} \left( \frac{\Sigma}{m} f_{Jm} t_{mrA} + \frac{\Sigma}{H} f_{JMB} t_{MrB} \right)^{2} + K = 0, \quad (B.4)$$

and B.3 becomes

$$g_{r} = \frac{1}{K} \sum_{J} \left( \sum_{m} f_{JmA} \cdot mrA - \sum_{M} f_{JMB} \cdot mrB \right)^{2}$$
(B.5)

Using LaGrange's system of undetermined multiplers, the minimum can be obtained when

$$\frac{\partial g_r}{\partial t_{mrA}} + \beta_r \frac{\partial \theta_r}{\partial t_{mrA}} = 0.$$
 (B.6A)

Substituting the indicated partial derivatives into B.6A,

When the equation is expanded, and terms are regrouped,

$$(1+ \kappa \beta_r) \sum_{J} f_{JmA} (\sum_{m} f_{JmA} t_{mrA}) = (1 - \kappa \beta_r) \sum_{J} f_{JmA} (\sum_{m} f_{JMB} t_{mrB}).$$
(B.8A)

Writing this equation in matrix form,

$$(1 + K\beta_r) F_{JmA}^i F_{JmA}^T T_{mrA} = (1 - K\alpha_r) F_{JmA}^i F_{JMB}^T T_{mrB}.$$
 (B.9A)

Define 
$$\phi_{\mathbf{r}} = \frac{(1 + K\beta_{\mathbf{r}})}{(1 - K\beta_{\mathbf{r}})}$$
 (B.10A)

Then

(B.11A

Similarly, finding derivatives with respect to the leads to

(BALL)

(Equation B.11B can be written from B.11A by interchanging  $\underline{\mathbf{m}}$ 's end  $\underline{\mathbf{M}}$ 's, and  $\underline{\mathbf{A}}$ 's and  $\underline{\mathbf{B}}$ 's.)

The meaning of  $\phi_r$  can be obtained by premultiplying equation B.llA by  $T'_{mrA}$  which gives

Substituting from equations B.1A and B.1B yields

$$F_{JrA}F_{JrB} = \phi_r F_{JrA}F_{JrA}. \tag{B.15A}$$

Writing equation B.13 in summational rotation,

$$\sum_{J} f_{JTA} f_{JTB} = \oint_{TJ} f_{JTA}^{2}; \qquad (3.14A)$$

02.

$$\phi_{\mathbf{r}} = \frac{\sum_{j}^{r} f_{jrA}^{f} f_{jrB}}{\sum_{j}^{r} f_{jrA}^{f}}$$
 (2.15A)

Similarly from equation B.11B it can be shown that

$$\phi_{\mathbf{r}} = \frac{\sum_{J} f_{JrA}^{T} J_{JrB}}{\sum_{J} f_{JrB}^{Z}}, \qquad (3.158)$$

Noting that the numerators of equations B.15A and B.15B are equal, these equations solved simultaneously give

$$\sum_{J} f_{JrA}^2 = \sum_{J} f_{JrB}^2$$
 (3.16)

Therefore, the sum of squares of loadings in one study on a congruent factor is equal to the sum of squares of loadings in the other study on this factor.

In order to relate  $\frac{\phi_{r}}{r}$  and  $\frac{\phi_{r}}{r}$  equation IV.3 is expanded to

$$g_{r} = \frac{\sum_{J} f_{JrA}^{2} + \sum_{J} f_{JrB}^{2} - 2\sum_{J} f_{JrA} f_{JrB}}{\sum_{J} f_{JrA}^{2} + \sum_{J} f_{JrB}^{2} + 2\sum_{J} f_{JrA} f_{JrB}}.$$
 (B.17)

Substitutions from equations B.15A, B.15B, and B.16 yield

$$g_{\mathbf{r}} = \frac{1 - \phi_{\mathbf{r}}}{1 + \phi_{\mathbf{r}}}$$
 (B.18)

From equation B.18 it is to be noted that  $g_{r}$  is a minimum when  $g_{r}$  is a maximum. (Only the positive range of  $g_{r}$  need be considered.) Making use of equation B.16, equation B.15A may be rewritten:

$$\phi_{r} = \frac{\sum_{J} f_{JrA} f_{JrB}}{\sqrt{(\sum_{J} f_{JrA})(\sum_{J} f_{JrB}^{2})}},$$
 (B.19)

which is similar to the formula for the product-moment correlation between the loadings on factor r for studies A and B. The difference from the equation for a correlation is that no corrections are made in this case for means of the factor loadings. Thus,  $\phi_r$  might be called a coefficient of congruence. This is in contrast with  $g_r$  being called an index of congruence.

One consequence of equation B.19 is that  $f_{\mathbf{r}}$  can never be greater than unity. A value of unity would indicate perfect congruence of the factor in the two studies. Values of  $f_{\mathbf{r}}$  less than unity indicate various degrees of congruence down to no congruence at a value of zero. For practical purposes it may be desirable to set up some value of  $f_{\mathbf{r}}$  less than unity which will be regarded as acceptable for indicating the identity of the factors in the two studies. However, no guiding values have yet been developed, and it seems proper to delay specifying any minimally acceptable value of the coefficient of congruence until adequate experience in the application of the method has been gained.

In solving equations B.11A and B.11B it seems advisable to obtain the latent roots and latent vectors of the matrices  $(F_{JM}^*F_{JMA}^*)$  and  $(F_{JMB}^*F_{JMB}^*)$ . Let  $\Lambda_{mpA}$  be an orthogonal transformation containing the latent vectors and  $\beta_{pA}$  be a diagonal matrix containing the latent roots of  $(F_{JMA}^*F_{JMA}^*)$ . Then,

$$(\mathbf{F}_{\mathrm{JmA}}^{i}\mathbf{F}_{\mathrm{JmA}}) = \mathbf{A}_{\mathrm{mpA}}\,\boldsymbol{\beta}_{\mathrm{pA}}\,\mathbf{A}_{\mathrm{mpA}}^{i}, \tag{B.20A}$$

$$(F'_{JMB}F_{JMB}) = \Lambda_{MPB} \beta_{PB} \Lambda'_{MPB}$$
 (B.208)

See Tables 10 and 11.

The latent vectors are frequently called principal axes in factor analysis. Let the matrix of factor loadings on the principal axes be  $\mathbf{F}_{JpA}$ . Then,

$$F_{JrA} = F_{JmA} \Lambda_{mrA};$$
 (B.21A)

$$F_{JPB} = F_{JMB} A_{MPB}. ag{B.21B}$$

See Table 12.

Since  $\Lambda_{mpA}$  is an orthogonal matrix,

$$F_{JmA} = F_{JmA} \Lambda_{mmA}^{\prime};$$
 (B.22A)

Substitution of equation B.22A in equation B.20A and simplification yields

$$\beta_{1A} = F_{JDA}^{\dagger} F_{JDA} \tag{E.23A}$$

$$S_{PB} = F_{PB}^{\dagger} F_{PB} \qquad (B.23B)$$

Each latent root, located in the diagonal of  $\mathcal{F}_{pA}$ , is the sum of squares of the corresponding column of  $F_{JyA}$ . The principal axis with the smallest latent root has the property of being the factor in the space defined by  $F_{JmA}$  with the minimum sum of squares of loadings. The principal axis with the next to smallest latent root has the minimum sum of squares in the space

orthogonal to the principal axis with the smallest latent root. statements concerning minimum sum of squares of loadings can be made concerning the remaining principal axes, taking them in order from smallest to largest, in each case considering the space orthogonal to the preceeding axes. Since the latent roots are these sums of squares of the loadings on the principal axes, these roots represent indices of the extent to which the tests project into the dimensions represented by the principal axes. Whenever a latent root is zero, all tests must have zero loadings on that principal axis. As indicated in detail later, this condition of a zero latent root causes the solution to equations B.llA and B.LIB to be non-unique. A small, but non-zero latent root indicates small loadings. This condition is likely to occur when the overlap tests represent some of the factors in the study but have only small random loadings on the other factors. Such principal axes with small latent roots may be delegated to a non-congruent space of the study by limiting the congruent factors to the space defined by the principal axes with significant latent roots. Study B, the four principal axes with smallest latent roots were placed in the non-congruent space. No precise rule has been developed for dividing between those principal axes to be delegated to the non-congruent space and the exes to be used in determining the congruent space. It is important, however, to exclude from the congruent space those dimensions into which the overlap tests have small projections.

Consider, for the present, that no latent rectors have been dropped from A hard (or A HPB) and substitute equation E.22A into equation B.1A. Then:

$$F_{JrA} = F_{JrA} \Lambda'_{mrA} T_{mrA}; \qquad (B.24A)$$

$$F_{JrB} = F_{JPE} \Lambda^{\prime}_{MFB} T_{MrB}. \tag{B.24B}$$

Define:

CANDER A PROPERTY PROPERTY OF THE PARTY OF THE PROPERTY OF THE PARTY O

$$T_{prA} = \Lambda'_{mpA} T_{mrA};$$
 (B.25A)

$$T_{PrB} = \Lambda_{MPB} T_{MrB}. \tag{B.25B}$$

Then, from equation B.24A:

$$F_{JrA} = F_{JrA}^{T}_{prA};$$
 (B.26A)

When equation B.25A is solved for TmrA,

$$T_{mrA} = A_{mrA}T_{prA}; (B.27A)$$

Substitution of equations B.20A, B.22B, and B.23A into equation B.11A yields:

$$\Lambda_{mpA}F_{JpA}^{\dagger}F_{JpB}\Lambda_{MPB}^{\dagger}F_{MrB} = F_{r}\Lambda_{ppA}B_{pA}\Lambda_{mpA}^{\dagger}F_{mrA}.$$
 (B.28A)

Simplification by means of equation B.25 yields:

$$F_{JpA}F_{JPB}T_{prB} = \phi_r \beta_{pA}T_{prA}. \tag{B.29A}$$

Similarly:

$$F_{\text{JPB}}^{\prime}F_{\text{JPA}}^{\dagger}T_{\text{DPA}} = \oint_{\mathbf{r}} \beta_{\text{PB}}^{\dagger}T_{\text{PPB}}.$$
 (B.298)

The matrix product F; FJPB for the example is given in Table 13.

Consider now the case when one of the latent roots in one of the studies is zero. Without loss of generality of the development, the last latent root for study A can be taken as the particular zero latent root. Any other latent root could have been chosen, it is merely a matter of convenience. The last diagonal entry in  $\beta_{\rm pA}$  is then zero. Since  $\beta_{\rm pA}$  is a diagonal matrix, the product  $\beta_{\rm pA}$   $T_{\rm prA}$  in equation B.21A results in the rows of  $T_{\rm prA}$  being multiplied by the corresponding latent roots. The last entry in  $T_{\rm prA}$  is, thus, multiplied by zero. When a latent root is zero, however, the leadings on the corresponding principal axis are zero. The last row of  $F_{\rm JpA}$ , therefore, has zero entries and the last entry in the product  $F_{\rm JpA}^{\rm T} F_{\rm JpB}^{\rm T} F_{\rm PRB}$  is zero. Therefore,

$$0 = \phi_r \cdot 0 \cdot \epsilon_{qrA}$$

where  $t_{qrA}$  is used to denote the last entry in  $T_{prA}$ . This equation is true no matter what value is given to  $t_{qrA}$ . In equation B.29B, the last column of  $F_{JpA}$  is zero; and, therefore, the value of  $t_{qrA}$  does not affect the equation. The conclusion is then that the value of  $t_{qrA}$  is not determined

by equation B.29A and B.29B when the corresponding latent root is zero. Thus, in this case, the solution is not unique. It is reasonable, however, to assign a value of zero to  $t_{qrA}$ . The congruent factors, then, will not involve this principal axis which can, thus, be delegated to the non-congruent space.

Other principal axes can be delegated to the non-congruent space by defining the corresponding entry in  $T_{prA}$  (or  $T_{prB}$ ) to be zero. Whenever any of the entries in  $T_{prA}$  are defined as zero, these entries, the corresponding columns of  $F_{JPB}$  and  $A_{mpA}$  (or  $F_{JPB}$  and  $A_{MPB}$ ), and corresponding rows and columns of  $\beta_{pA}$  (or  $\beta_{pA}$ ) can be dropped without affecting equations B.26A, B.27A, or B.29A (or the corresponding B equations.) In the following developments, it will be considered that these matrices have been so reduced.

Consider the case when there are as many significant latent roots in one study as overlap tests.  $F_{JpA}$  (or  $F_{JPB}$ ) is square. This matrix then will possess an inverse. Equation B.29A can be solved for  $T_{prA}$ :

$$T_{prA} = F_{JpA}^{-1} F_{JPB}^{T}_{prB}. \tag{B.50A}$$

 $\phi_r$  is unity for all  $T_{PrB}$ . Any values can be assigned to  $T_{PrB}$  and a  $T_{prA}$  can be obtained. Thus, perfect congruence has been obtained as a mathematical necessity irrespective of the characteristics of the tests in the other study. No confidence can be placed in the observance of such a congruence of factors.

Since only those principal axes have been retained that have significant latent roots the \$\beta\$ matrices are non-singular and possess inverses. It is now assumed that there are more overlap tests than significant latent roots in either study. It is convenient to define the column vectors:

$$M_{rA} = \beta_{pA}^{\frac{1}{2}} T_{prA}; \qquad (B.51A)$$

$$M_{rB} = \beta \frac{2}{\rho_B} \eta_{PrB}; \tag{B.31B}$$

$$G = \beta_{pA}^{\frac{1}{2}} F_{jpA}^{j} F_{JPB}^{j} \beta_{PB}^{\frac{1}{2}}.$$
 (B.32)

The matrix G for the example is given in Table 14.

Solution of equation B.51A for TprA yields:

$$T_{prA} = \beta_{pA}^{-\frac{1}{2}} N_{rA}^{2}$$
 (B.85A)

$$T_{PrB} = \beta_{PB}^{\frac{1}{8}} M_{PB}. \tag{B.35B}$$

Substitution of these equations into equation B.29A yields:

$$G H_{rB} = \oint_{r} H_{rA}$$
 (B.34A)

$$G'M_{rA} = \phi_rM_{rB}$$
 (B.34B)

Solution of equation B.34 B for  $M_{\rm PB}$ , substitution in equation B.34A, and simplification yields:

$$GG'M_{rA} = \phi_{r}^{eM}MA. \qquad (B.35A)$$

Similarly:

$$G \cdot GM_{pB} = g_p^R M_{pB^*}$$
 (B.35B)

Define:

$$H_A = GG^{\dagger};$$
 (B.36A)

$$E_{B} = G'G. (B.36B)$$

See Table 14.

Then:

$$H_{\mathbf{A}}^{H}\mathbf{r}_{\mathbf{A}} = \mathcal{G}_{\mathbf{r}}^{2}\mathbf{H}_{\mathbf{A}}; \tag{9.37A}$$

$$H_{B}H_{rB} = \phi^{2}H_{rB}. \tag{B.37B}$$

The matrices  $H_A$  and  $H_B$  are Gramian with latent roots  $\mathcal{G}_{\mathbf{r}}^2$  and latent vectors  $\Lambda_{\mathbf{r}A}$  and  $\Lambda_{\mathbf{r}B}$ . The non-vanishing latent roots are identical for the two matrices. For each latent root there is a latent vector  $\Lambda_{\mathbf{r}A}$  for  $H_A$  and a latent vector  $\Lambda_{\mathbf{r}B}$  for  $H_B$ . Table 15 gives the results for the example

When  $M_{rA}$  and  $M_{rB}$  are defined as proportional to the latent vectors, equations 5.37A and B.37B are solved since these equations are in a standard form for latent roots and vectors. Then:

$$M_{rA} = A_{rA} d_{rI} ag{B.58A}$$

$$\mathbf{M}_{\mathbf{r}\mathbf{B}} = \mathbf{A}_{\mathbf{r}\mathbf{B}} \mathbf{d}_{\mathbf{r}}. \tag{B.588}$$

Note that the same constant of proportionality,  $d_{\mathbf{r}}$ , is used for both studies.

Substitution of equations B.38A and B.38B into equations B.33A and B.33B yields:

$$T_{prA} = \beta_{pA}^{\dagger} \Lambda_{rA} d_{r}; \qquad (B.59A)$$

$$\gamma_{PrB} = \beta_{PB}^{\frac{1}{4}} \Lambda_{rB}^{\dagger} \alpha_{r'} \qquad (B.39B)$$

Substitution of these equations in B.27A and B.27B jields:

$$T_{mrA} = \Lambda_{mpA} B_{pA}^{\frac{1}{2}} \Lambda_{rA} d_{r}; \qquad (B.40A)$$

$$T_{MrB} = \Lambda_{MFB} \beta_{PB}^{-\frac{1}{2}} \Lambda_{rB} d_{r}.$$
 (B.408)

It is convenient to define:

$$T_{\text{mrA}} = \Lambda_{\text{mpA}} \beta_{\text{rA}}^{\frac{1}{2}} \Lambda_{\text{rA}}$$
(B.41A)

$$T_{\text{MrB}} = {}^{\Lambda}_{\text{MPB}} {}^{R}_{\text{FB}} {}^{\Lambda}_{\text{FB}}$$
 (B.418)

See Table 15.

Then:

$$T_{mrA} = T_{mrA} d_r$$
 (B.42A)

The constant,  $d_{r}$ , (one for each congruent factor) can be determined such that the average sum of squares of the entries in  $T_{mrA}$  and  $T_{MrB}$  is unity:

$$\frac{1}{2} \left( \sum_{\mathbf{H}} \mathbf{t}_{\mathbf{mrA}}^{2} + \sum_{\mathbf{H}} \mathbf{t}_{\mathbf{HrB}}^{2} \right) = 1.$$
 (B.43)

This is similar to the usual practice in factor analysis of making the sums of squares of entries in a factor transformation vector unity. In this case, the transformation vectors for the two studies can not be normalized separately. It is, then, reasonable to normalize the two vectors on the average. In order to accomplish this step:

$$d_{\mathbf{r}} = \frac{1}{\sqrt{\frac{1}{2}(\frac{\Sigma T^2}{n} + \frac{\Sigma T^2}{M} + \frac{\Sigma T^2}{M} + \frac{\Sigma T^2}{M} + \frac{\Sigma T^2}{M} + \frac{1}{M} + \frac{1}{M}$$

where the r's are the entries in the T vectors. Table 16 contains the  $d_r$ 's for the example. The transformation matrices  $T_{mrA}$  and  $T_{MrB}$  are in Table 5.

When the transformations to congruent factors,  $T_{mrA}$  and  $T_{MrB}$  have been determined, the matrices  $F_{JrA}$  and  $F_{JrB}$  of loadings on the congruent factors can be obtained by equations B.1A and B.1B. Table 4 contains these matrices of loadings on congruent factors for the example.

The computing procedure given in Appendix C is based on the foregoing equations. A simplification of steps was obtained by defining

$$T_{mpA} = \Lambda_{mpA} S_{pA}^{\frac{1}{2}}; \qquad (B.45A)$$

Only those principal axes with significant latent roots are included in these equations. Substitution of equations B.21A, B.21B, B.45A, and B.45B into equation B.32 yields:

$$G = T_{mpA}^{\prime} (F_{jmA}^{\prime} F_{jBB}) T_{MPB}. \tag{B.46}$$

With the definitions of equations B.45A and B.45B, equations B.41A and B.41B become:

$$T_{mrA} = T_{mpA} \wedge rA$$
 (B.47A)

$$T_{MrB} = T_{MPB} \Lambda_{rB}$$
 (B.478)

When the congruent factors and coefficients of congruence have been determined, one or more of these factors may be judged not to be sufficiently similar to be continued in the congruent space. This is indicated by a low coefficient of congruence (or a high index of congruence). In the example, factor F, with a coefficient of congruence of .459717, was eliminated. The remaining factors then define the congruent space between the two studies.

After the congruent space between two studies has been determined by the congruent factors, rotation within this space is possible. Consider that it is desired to find the loadings on a set of factors s which are defined as linear combinations of the congruent factors. Let the coefficients for these linear combinations be included in a matrix  $T_{\rm rs}$ . There will be a column for each factor s and a row for each congruent factor r. In the example, after an inspection of the loadings on the congruent space, it was decided to:

- Perime five vectors as the sums of test vectors in the congruent space for five sets of tests indicated in Table 17.
   The sums of loadings on the congruent factors for the tests in each group are also given in Table 17.
- 2. Define a set of five factors so that each factor would have zero loadings for four of the summation vectors. Each factor would be defined by the summation vector with a non-zero loading. This could be accomplished by computing the inverse of the matrix of loadings of the summation vectors on the congruent factors. The matrix Trs in Table 18 is this inverse.

The factors a are related to the reference axes by the equations:

$$T_{\text{msA}} = T_{\text{mrA}} T_{\text{rs}}$$
 (8.48A  
 $T_{\text{MsB}} = T_{\text{MrB}} T_{\text{rs}}$  (8.48B

See Teble 18.

The transformation vectors in  $T_{\rm msA}$  and  $T_{\rm MsB}$  are normalized on the average between the two studies just as the congruent factors were normalized on the average

between studies in equations B.42 - B.44. The equations for this step are:

$$d_{B} \sqrt{\frac{1}{2} \left(\frac{\Sigma T^{2}}{meA} + \frac{\Sigma T^{2}}{meB}\right)_{1}}$$
 (B.49)

$$T_{msA} = T_{mrA} d_{s}; (B.50A)$$

See Table 18.

Loadings on these factors are found by:

$$\mathbf{F}_{JBB} = \mathbf{F}_{JMB} \mathbf{T}_{MSB}$$
 (B.518)

See Table 19.

An alternative method is to define:

$$T_{rs} * T_{rs} d_s. ag{B.52}$$

See Table 18.

Then:

$$\mathbf{r}_{JsA} = \mathbf{r}_{JrA} \mathbf{r}_{rs};$$
 (B.53A)

$$F_{JaB} = F_{JrB} T_{rs}$$
 (B.538)

A point to note is that since the leadings in  $F_{JrA}$  and  $F_{JrB}$  are similar due to the solution to congruence, the leadings in matrices  $F_{JsA}$  and  $F_{JsB}$  must also be similar.

If it is desired to rotate the exes to a new position z, a similar cycle to the preceeding is taken. In equations B.48A - B.53B, substitutions of s for r and of z for s are made. In the example, a rotation as indicated in Table 20 was decided on. The results of this rotation are given in Tables 7-9.

The loadings of the non-overlap tests may be obtained at any time in this rotational procedure by extending equations B.51A and B.51B to these areas of the reference factor matrices.

Consider now the non-congruent factors. There will be a set for each study. Let u and U designate the non-congruent factors in studies A and B. The transformations to these factors can be defined by the equations:

$$T_{mrA}^{\dagger} A_{muA} = 0;$$
 (B.54A)

$$T_{MrB}^* \Lambda_{MUB} = 0; (B.54B)$$

$$\Lambda_{\text{KIB}} \Lambda_{\text{MUB}} = 1.$$
 (B.55B)

These transformations are sections of orthogonal transformations with as many columns as there are non-congruent factors in the studies. A computing procedure for solution of these equations is given in Appendix C. Results for the example are given in Table 6.

Table 10

matrices (Fire<sup>F</sup>jea<sup>A</sup>) and (Fire<sup>F</sup>jeb) study a matrix (Fjea<sup>F</sup>jea<sup>A</sup>)

	74	3.88.9.9.9.9.9.9.9.9.9.9.9.9.9.9.9.9.9.9
	*	1.094 1.00 1.00 1.00 1.00 1.00 1.00 1.00 1.0
Factors	Śν	1.180 134 205 205 706 714
Reference 1	111	.501 1115 1265 1265 1265
	11	601 44.5 44.1 46.0 688
	Ψł	3.177 
and the state of t	Factors	**************************************

# STUDY B MATRIX (FINEFURS)

·Ħ	<b>89</b>
Ħ	27. 20. 20. 20. 20. 20. 20. 20. 20. 20. 20
×	200 200 200 200 200 200 200 200 200 200
Ħ	647. 145. 145. 125. 127. 100.1 100.0 100.0
VIII	505. 100. 100. 100. 100. 100. 100. 100.
e Factors	200. - 008 - 00. - 140. - 140. - 100. - 100.
Reference 1	20.1 100.1 100.1 100.1 100.1 100.1 100.1 100.1 100.1 100.1 100.1 1
>	ૠ૽૽ૼઌ૽૽ૹ૽ૹ૽ૹ૽ૹ૽૽ ૹ૽૽ૹ૽ૹ૽ૹ૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽ ૡઌ૽૽ઌ૽૽ૹ૽ૹ૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽ૹ૽૽
A	78. 25. 26. 26. 27. 27. 27. 27. 27. 27. 27. 27.
Ħ	155 155 155 157 272 272 273 273 274 203 214 215 215 215
Ħ	
₩	1.050 1.050 1.050 1.030
Reference	r ri ry ry riii RX XX

Heference	Principal Axes									
<b>Factors</b>	<b>p-1</b>	p-2	p-3	<b>p-4</b>	p-5	p-6				
1	.8208	.4129	2910	2189	.1289	-,0800				
11	.1524	.3925	.8975	.1030	0617	0561				
111	-1258	.1799	1373	.5118	3608	.7355				
iv	.3689	4269	.0059	6032	2549	5026				
₩	.3431	6006	.2603	4972	5147	.3282				
<b>V1</b>	-1826	-,3164	.1529	.2618	.8280	.2987				
$\boldsymbol{\beta}_{\mathbf{pk}}$	4.4565	.7939	.3062	.2104	.0864	.0378				
$\beta_{pA}^{\frac{1}{2}}$	2.1110	.8 <del>9</del> 10	.5534	.4587	.2939	.1945				
Bon	.4757	1.1223	1.8072	2.1800	3,4029	5.1407				

#### STUDY B MATRIX A

Reference						Princi	pal Axes	ı		-		
<b>Factors</b>	P-1	P-2	P-3	P-4	P-5	P-6	P-7	<b>P-8</b>	P-9	P-10	P-11	P-12
I	.3713	١٠;91	1522	1588	0168	.5476	4832	2697	0043	0592	0244	0599
II	.0109	.1588	.1100	0937	.0187	.2726	,1817	0002	.1890	2934	.8158	2423
III	.1119	2074	<b>.049</b> 4	1777	1454	4346	3493	.3719	- 5234	c\76	.2507	3179
IY	.5615	.6085	0852	4637	.1194	.0223	0002	.1536	0288	1110	1946	.0522
٧	.2451	1450	.5064	.0837	6084	.0372	.1856	.1203	.1235	4015	2307	.0494
AI	.3737	5286	1354	2626	.1696	2437	.5881	.0102	.1315	.1839	.0313	.0437
VII	0055	0557	.0941	10':6	+.0001	-,1405	-,2216	.0604	0064	وأبانه	.3262	.8911
VIII	.2455	.1350	6732	.3787	5151	1317	.0125	0162	.0861	.0407	.1676	.0511
IX	.3832	3777	وجشاه	.288)	0416	0156	~.0035	31,25	2141	•5799	.1935	0095
X	.0101	02.62	0468	0275	.0834	3657	.0567	7395	2189	4920	0253	0163
XI	.5452	0697	.Cl:17	6249	.5391	0707	0019	2602	.0662	3271	0238	.0006
XII	.0432	.0052	.1372	1166	0050	4468	4128	1224	.7532	.0981	.0179	1858
$oldsymbol{eta}_{ exttt{PB}}$	5.5%or	.9522	.3425	.2638	.1059	.0675	.0615	.0357	.0123	•0000	.0005	o100.
8 22 YB	2.4552	.9158	.5853	.5136	.3254	.2598	.2480	,1890	••••	••••	••••	
β <sub>PB</sub> <sup>1</sup> / <sub>2</sub>	.4110	1.0248	1.7087	1.9469	3.0729	3.8496	4.0330	5.2918	••••	*****	****	• • • • • •

e. State of all the second of the second state of the second of the seco

PROPERTY CONTROL OF THE PROPERTY OF THE PROPER

FACTOR LOADINGS ON PRINCIPAL AXES

## STUDY A MATRIX PJEA

	9-4	9050	.0396	1412	.0185	.0524	88.	.0208	0246	0690	0539
	\$-\$	9200.	9650.	8720.	.1320	888	.0713	2020.	2020	1289	888.
Axes	7 0	6001	+161·-	.1055	3081.	.0585	-1042	1854	1248	.2235	.075±
Princing	6-q	.0200	1718	1358	0456	1699	.0621	0+5t	.0273	2188	1901
	e e	84:68					_		_		
,	p-1	5037	.6581	7418	.6621	6639	18131	1089	.5373	7177.	5948
	study B	м	56	233	ଝ୍ଯ	55	N	Ħ	33	27	8
Moor Code	Study A	97	o	, tr	12,	-4	121	7	15	ì	16+17

## STUDY B MATRIX PAPS

-56-

	त-र		
	F-11		
	<b>P-1</b> 0	.0013 .0024 .0020 .0010 .0010 .0000 .0000 .0000	
	<b>8-9</b>	2010. 2010. 2010. 2010. 2010. 2010. 2010. 2010.	
	87	.0536 .0120 .0991 .0005 .0005 .0009 .0085 .0636	
•	1-d	0590 0542 0189 0457 0257 0126	
	<b>P-6</b>		
	P-5	0542 0346 0346 0368 0369 0361 0316 	
	P-4	-3019 4034 6035 6031 2603 -0573 1704 1704 1704 1704 1704 1704 1704 1704	
	ቴ የ		
	4		
	т <del>-</del> й	6203. 6203. 6779. 6779. 607. 607. 607. 607.	
	Study B	1,89,88,88,828	
	Test Code N Study A S	83 0 2 7 14 21 14	

### Table 13 THE MATRIX (F; FJPB)

Study A		Study B Principal Axes							
Principal Axes	. P-1	P-2	P-3	P-4	P-5	<b>P-</b> 6	P-7	<b>P-8</b>	
p-1	5.1251	0021	.0028	0133	0218	0065	.0092	0174	
p-2	0015	8259	1292	.0029	.0051	0065	.0407	.0041	
p-3	0491	.0066	.0925	1101	0761	0092	.069 <b>9</b>	.0207	
p-4	0059	0648	.1902	.0831	.0451	0483	.0159	0190	
p-5	0056	0496	.0544	.0619	0609	.0182	0293	.0019	
p-6	.0003	0094	0115	.0035	.0211	0139	0218	.0100	

Table 14

#### THE MATRICES G, HA, HB

#### MATRIX G

Study A Principal Axes	P-1	P-2	Study P-3	B Princ P-4	ipal Axe P-5	P-6	P-7	P-8
p-1	.9977	0010	.0023	0123	0317	0119	.0176	0436
p-2	0007	9499	2478	.0063	.0176	0281	.1842	.0243
p-3	0365	.0122	.2856	3874	4226	0640	.5094	.1980
p-4	0053	1448	.7085	.3527	.3021	4053	.1398	2192
p-5	0078	1730	.3163	.4101	6368	.2384	4021	.0342
p-6	.0006	0495	1010	.0350	.3333	2751	4520	.2721

#### MATRIX H

Study A						
Principal Axes	p-1	p-2	ly A Prin p-3	p-4	p-5	p-6.
p-1	.9989	.0016	0165	-,0006	.0031	0271
D-5	.0016	.9994	.0082	.0013	0026	.0092
p-3	0165	.0032	.7145	-,0098	0146	3426
7-4	0006	.0013	0098	.9705	.0411	.0373
p-5	.0031	0026	0146	.0111	.9234	0958
n-6	0271	.0092	3426	.0373	0958	.4790

#### matrix h<sub>B</sub>

Study B		Study B Principal Axes									
Principal Axes	P-1	P-2	P-3	P-4	P-5	P-6	P-7	<b>P-8</b>			
P-1 P-2	.9968 .0013	.0013	0142	0032 1314	0126	0094 .0570	.0010 0971 .1174	0497 0083 1215			
P-3 P-4 P-5	0142 0032 0126	.0866 1344 .0281	.7552 .2638 1462	.2638 .4441 .0213	1462 .0213 .7878	1953 0300 3390	3278 0649	1298 0792			
P-6 P-7 P-8	0094 .0010 0497	.0570 0971 0083	1953 .1174 1215	0300 3278 1298		.3018 0662 .0093	0662 .6793 0628	.0093 0628 .1650			

Table 15

### latent vectors and roots for matrices $\mathbf{H}_{\mathbf{A}}$ and $\mathbf{H}_{\mathbf{B}}$ study a matrix $\mathbf{A}_{\mathbf{r}\mathbf{A}}$

Study A	Congruent Factors									
Principal Axes	A	В	C	D	E	T				
p-1	.8341	3459	.0031	.3569	2356	.0408				
p-2	.4218	.8889	0367	0073	.1741	0151				
p-3	.1541	1579	.5028	1999	.5797	.5679				
<b>p-</b> 4	2803	.1481	.4512	.8318	.0492	0397				
p-5		1198	7019	.3595	.5885	.1250				
p-6	1494	.1702	2225	.1071	4789	.8114				

#### STUDY B MATRIX A TB

Study B	Congruent Factors									
Principal Axes	A	В	C	a	E	F				
P-1	.8273	3398	0131	.3571	2627	.0419				
P-2	3443	8552	.1154	1842	2444	0740				
P-3	2548	2164	.2914	.6393	.3930	.2030				
P-4	1880	.0801	3540	.5188	.0214	3295				
P-5	1925	.2711	.3160	.1314	7574	1334				
P-6	.1132	1462	3419	2728	.2139	4546				
P-7	.2154	.0692	.7425	1740	.3113	2878				
P-8	.0238	.0152	0908	1967	.0084	.7318				

LATERT ROOTS \$2

A	В	C	Œ	E	F
1.00173	.99965	.87694	.99517	.99092	.22109

Congruent Factors

Table 16  ${\it MATRICES}~T_{\it MTA},~T_{\it MTB},~\cdot {\it d}_{\it T}$  STUDY A MATRIX  $T_{\it MTA}$ 

Reference						
Factors	A	B	C	D	E	y
1.	.6161	.1772	7118	0428	.1160	5498
ii	.4847	.1199	1.1126	2212	1.0255	.6484
iii	7309	1.1538	.3917	•9609	-2.6011	2.7287
iv	0019	6290	1.8005	•5690	.6730	-2.2366
v	.0191	4748	.1492	-1.1368	-1.3746	1.5622
<b>vi</b>	5394	3809	-1.9095	1,6303	1.0289	1.7409
E mrA	1.440049	2.139747	8.808218	5.248043	11.231917	18.642088

#### STUDY B MATRIX T MTB

Reference			Congruent	Factors		
Factors	A	В.	č	D	E	r
r	.2039	1054	-2.0761	1493	1649	-1.3456
II	.1989	2836	.3446	4026	.4379	- • 6334
III	2052	.1760	6718	.1857	3759	2.8145
ΊΛ	.1412	5532	.3583	7290	5427	.7386
Λ	.4332	5h76	.0317	.1682	2.0401	.5692
VI	.7743	.8637	2.2904	3573	.1337	1184
VII	2307	.0215	395 <b>3</b>	.2472	3179	.8407
VIII	.4477	1978	8629	4035	.6084	1180
IX	2553	3853	.1715	1.1550	.2936	-1.3153
X	2321	.2097	1.0874	1.0633	5003	-2.3255
XI	5017	.5791	2039	.8262	-1.3824	.6422
XII	5652	.0435	4487	-9147	-,8045	,9094
ET 2 MrB	1.874244	2.043654	12.611188	5.114092	8.200618	20.114876
2(ET2 + ET2)	1.657146	2.091700	10.709703	5.181068	9.716268	19.378482
d <sub>r</sub>	.776819	.691433	.305570	.439329	.320812	.227164

Table 17
SUMMATION VECTORS FOR FIRST ROTATION IN CONCRUENT SPACE

Summation		used Test	Sum of Loadings on Ucngruent Factors									
Vector	From Study	Number	A	B	C C	D	2					
8.	A A B	18 9 1 · 26	1.3562	.4186	3866	1810	0415					
ъ	A B	4 22	.0133	7811	.0071	.1880	0156					
c	A B	15 33	.3476	5268	.2662	3428	2937					
đ	A B	11 27	.3006	.5138	.1349	.4511	4279					
e	A B	16+17 30	.8516	.2690	.3318	.0193	.3379					

Ø		
Ŋ		
Š		
73		
37		
44		
R		
13		
葡		
3		
Ė		
И		
7		
2		
22		
37		
à		
1		
5		
à		
ŧ		
ş		
3		
THE PROPERTY OF THE PROPERTY O		
1		
i		
- 1		
- 1		
- 1		

TORREST CHARGE CHARGES CHARGES THE STATES TO THE STATES THE STATES TO THE STATES OF TH

				•					•	-4	68-													
			•	.0578 .0578 .5981		•	0173	5 S	<b>8</b>	. 0571			•	1144	272	6110 9013	6306	.1629	525	XISS ZIZS				
			Petore	8832 7272 7272 7273 7273 7273 7273 7273		Pectors	0342	4721-	80	1000		Pactors d	-1783	8	-1060	200	8	8203	, 35 E	•		•		
		MATRIX T.	Motated Congruent Pastory b c c d	2500 2500 2400 2600 2600 2600 2600 2600 2600 26	MANNET THE	btated Congruent Pactors b c d	2450-	98.	5750			MATERIX TAEB	Rotated Congruent Pactors b c d	2122	i.	4247	9	88	14.7	हुई हैं				
		₹	Notated C	2179 27/5 27/5 7882 1516	*	Notested C	1583°		3414	8		3	Rotated C	1918	8	0592	-5373	9.0°	2	2056	•			
	PACK		•	4679 -1.0099 -1.0099 -1.567		•	. 4368	2085	-6100	2558			•	7338	127	8.50 6.60 6.60 6.60 6.60 6.60 6.60 6.60 6	45.	9.7 2.7	2188	895	}			
	TRANSFORMATION MATRICES FOR FIEST BOLATION COMPRISE SPACE	•	Congruent Factors	<b>≪</b> #UA#		Reference	1**	ដដូ	÷	<b>&gt;</b> 7			Reference Factors	М	誯	A P	IA	11 Y	<b>"</b>	1#H				
Teble 18	OR FEET ROTH		•			•	+1.co-	. 1880 1880 1880	.8375	-:2775-	2.155620		v	6991	.3603	1235	\$66	28582	2012	4569 4319	2.867015	2.511518	.6310	
	N MATRICES #	unix I.	MATRIX Fra	ictore 4	. 1286 . 2471 . 3437 . 8535 . 956		se tors d	66HO	1962	ă,	0262	1.828451		actors d	2351-5	#17:-	1513	87.12	1257	0185	. 1518 1518 1964.	2.243561	2.038511	.700j
	PANSICHATIO			Notated Congruent Factors b	.1578 .2559 .1.0145	жыты тыл	Rotated Congruent Factors b	1536	8791.	.0931	1.62%	5.420938	MATRIX T <sub>NSB</sub>	Rotated Congruent Factors b	1691	1886	5738 7418	8,50	1430	5537	1641.	1.358819	2.639904	.6155
		Ž	Fotated b	. 9557 1.2955 1.2955 1.2955	×	Rotated	3581.	13.23	55.	1.3765	2.918422	2	Rotated b	5112	0072	.0962	6119	.0915	(g) (g)	.1350	2.349342	2.63388	6162	
			ď	.4068 4682 9685 1251		ವ	4150	-2000	158	.1505 545 545	.731024		4	.7038	.1318	885 845 845	4116	.0379	505.	350.	1.108795	.919910	1.0426	
			Congruent Factors	<b>≪</b> ₽0₽¥		*Reference Factors	Ħ	###	Ţ,	÷ 74	A Tre		Reference	러}	1日	A P	IA	IIV TITO	ä,	*##	E Y THE	2(ET2 4ET2)	જ	

Test Cod	le Numbers	Loadings on Rotated Congruent Factors - s								
Study A	Study B	<u></u>	<u>p</u>	<u>e</u>	<u>d</u>	<u>e</u>				
18	1	19	04	.02	.03	.03				
9 5	26	.32	.03	01	03	02				
5 ·	23	.02	.52	05	.03	.09				
12	28	.06	.30	- ,21	.13	.04				
4	55	03	.30	.02	.01	.02				
13	32	.19	.52	.01	01	03				
14	34	.23	.06	,06	03	<b>05</b>				
15	33	.00	.00	.31	.00	.00				
11	27	01	.00	.00	.35	.00				
16+17	30	.00	•00	.00	.00	.32				

#### STUDY B MATRIX F<sub>JSB</sub>

Test Cod	c Numbers	Loadings on Rotated Congruent Factors - s								
Study A	Study B	8	<u>b</u>	c	d	<u>e</u>				
30	1	.18	04	.03	.03	.03				
9	26	.36	.04	03	04	04				
9 5	23	.01	.32	04	.03	.10				
12	28	•04	.29	20	.14	.05				
4	55	.03	.32	02	01	02				
13	32	.13	.20	.04	.00	.02				
14	314	.23	.06	.05	03	~.05				
15	33	.00	.00	.31	.00	.00				
11	27	.01	.00	.00	.35	.00				
16+17	30	.00	.00	.00	.00	.31				

#### APPENDIX C

#### COMPUTING PROCEDURE FOR SYNTHESIS OF

#### FACTOR ANALYSIS STUDIES: .

The following set of notes give detailed computing procedures to implement application of the method presented in this report for synthesis of factor analysis studies. For convenience these notes are divided into five sections:

- 1. Congruent factor computations.
- 2. Rotation of axes in the congruent space.
- 3. Determination of non-congruent axes.
- 4. Determination of latent roots and vectors.
- 5. Notes on matrix computations.

Section 1 contains the basic elements of the method for synthesis of factorial studies. Section 2 and 3 pertain to subsequent steps. Sections 4 and 5 are included to facilitate the computations of the preceding three sections.

While it will be assumed that the work will be under the direction of a person competent in factor analysis, only a minimum knowledge will be assumed for the person doing the computations. It will be assumed that the person performing the computations is trained in the operation of a calculating machine and has some knowledge of statistical computations. No knowledge of matrix algebra nor of matrix computations will be assumed. Section 5 of these notes is intended to sumply the limited instruction necessary concerning matrices. While specific references will be made in the other sections to relevant portions of section 5, it would be advisable that the person doing the computations to become familiar with the contents of section 5.

#### 1. Congruent Factor Computations

The computational procedure for determining congruent factors will be illustrated by a fictitious example. Table 21 gives the factor matrices  $F_{jmA}$  and  $F_{jMB}$  and standard deviations  $\sigma_{jA}$  and  $\sigma_{jB}$  for six overlap tests in Studies A and B. Study A has three factors, Study B has four factors. Each row of each factor matrix has been summed with the sum being recorded in the  $\Sigma$  column.

#### s. Equalize Units of Measurement

A) Compute the constants  $d_{jA} = \sigma_{jA} / \frac{1}{2} (\sigma_{jA} + \sigma_{jB})$ , and  $d_{jB} = \sigma_{jB} / \frac{1}{2} (\sigma_{jA} + \sigma_{jB})$ . In order to facilitate computations a worksheet was set up as shown in Table 22.

The operations in the successive columns are as follows:

- 1) Enter the  $\sigma_{jA}$  for each test in Study A in column 1. Total the entries in column 1 and enter results in row  $\Sigma$  column 1.
- 2) Enter the  $\sigma_{jB}$  for each test in Study B in column 2. Total the entries in column 2 and enter the results in row  $\Sigma$  column 2.
- 3) Sum the  $\sigma_{jA}$  and  $\sigma_{jB}$  for each test and divide the sum by 2. Record the result in the third column designated  $\frac{1}{2}(\sigma_{jA} + \sigma_{jB})$ . In order to check the entries in the third column, total the entries in the E cells of the first and second column and divide the sum by 2. Enter the result in the Ch cell of the third column. Total the entries in the third column and enter in the E cell. The Ch and E entries should agree.
- 4) Divide the  $\sigma_{jA}$  and  $\sigma_{jB}$  in columns 1 and 2 for each test by the  $\frac{1}{2}(\sigma_{jA} + \sigma_{jB})$  in column 3 for the test and record in column 4 and 5,  $d_{AA}$  and  $d_{AB}$ .

- 5) Sum the d<sub>jA</sub> and d<sub>jB</sub> for each test and enter the total in column 6. The sum should equal 2.00.
- B) Compute the Reference Factorial Matrices for Tests with Adjusted Units of Measurement,  $F_{JinA}$  for Study A and  $F_{JMB}$  for Study B. See Table 23.
  - 1) To obtain the matrix  $F_{\text{JmA}}$ :
    - a) Multiply each entry in the first row of matrix F<sub>jmA</sub>, Table 21, by the constant d<sub>jA</sub> for the first test in column 4 of Table 22:

 $.35 \times .89 = .32$   $-.34 \times .89 = -.30$  $.48 \times .89 = .43$ 

b) In order to check the first row of F<sub>JmA</sub> multiply the sum of the first row of F<sub>JmA</sub> by the constant d<sub>JA</sub> for the first test in column 4 of Table 22:

.49 x .89 \* .44.

Record the result in the Ch column of  $F_{JmA}$ . Sum the entries in the first row of  $F_{JmA}$  and record in the  $\Sigma$  column. The entries in the Ch and  $\Sigma$  columns should agree within  $\frac{1}{2}$  of the last decimal place carried.

- c) Compute and check the entries in each row of  $F_{JmA}$  in the same way, using the corresponding row of  $F_{JmA}$  in Table 21 and the corresponding constant  $d_{JA}$  from column 4 of Table 22.
- d) Obtain the sum of each column of  $F_{JmA}$ .
- 2) Compute the matrix  $F_{JMB}$  following the same procedure as outlined for the computation of  $F_{JmA}$  by using the rows of  $F_{JMB}$  and the constants  $d_{JB}$ .
- b. Compute the principal exes for the tests in each study.
  - A) Compute the matrix product  $F'_{JmA}F_{JmA}$  of Table 24 (An explicit following of the formula  $F'_{JmA}F_{JmA}$  would involve (1) recording of  $F'_{JmA}$ , the transpose of matrix  $F_{JmA}$ , see Section 5,

Paragraph a, Point 6; and (2) matrix multiplication of matrices  $F_{,lmA}$  and  $F_{,lmA}$ , see Section 5, Paragraph c. It is unnecessary to record the transpose matrix  $F_{,lmA}^*$ , the same results may be obtained by multiplying each column of  $F_{,lmA}$  by every column of  $F_{,lmA}$ . The rows of the transpose matrix are implied by their equivalents, the columns of  $F_{,lmA}$ .)

## 1) First row of FinaFJmA

- a) Compute the sum of squares of the entries in the first column in  $F_{JmA}$  and enter the result in the first coll of the first row of  $F_{JmA}^{1}F_{JmA}$ .
- b) Compute the sum of products between the entries in the first column and in each other column of F<sub>JmA</sub> and enter the result in the corresponding cell of the first row of F'<sub>JmA</sub>F<sub>JmA</sub>. For example: the sum of products between the entries in the first column and second column of F<sub>JmA</sub> is entered in the first row, second cell of F'<sub>JmA</sub>F<sub>JmA</sub>.
- c) Compute the sum of products between the entries in first column and in the  $\Sigma$  column of  $F_{JmA}$  and enter in the Ch cell of the first row of  $F_{JmA}^*F_{JmA}$ .
- d) Sum the entires in the first row of  $F_{JMA}^*F_{JMA}$  (exclusive of the entry in the Ch cell) and enter the total in the  $\Sigma$  cell of the first row of  $F_{JMA}^*F_{JMA}^*$ . This entry should agree to within  ${}^{\frac{1}{2}}$  2 of the last deciral place carried.

# 2) Second row of FinaFJmA

Compute the second row of  $F_{JmA}F_{JmA}$  by using the second column of  $F_{JmA}$  and reporting preceding steps a-d. The sum of squares of the second column will be entered in the second call of the second row.

3) Remaining rows of FinAFJmA

Compute the remaining rows of  $F_{JmA}^{1}F_{JmA}^{2}$  by using the corresponding columns of  $F_{JmA}^{2}$  and repeating steps a.d. In each case, the sum of squares of the

columns of  $F_{JmA}$  will be the diagonal of  $F_{JmA}^{*}F_{JmA}^{*}$ . Note that  $F_{JmA}^{*}F_{JmA}^{*}$  is symmetric and once a row is computed and checked it may be copies into the corresponding column.

B) Compute the matrix product F; MBF, MB for Study B following the procedure as outlined in the preceding steps for

FjmAFJmA\*

- c) Solve for latent roots and vectors  $\Lambda_{mpA}$ ,  $\beta_A$  and  $\Lambda_{MPB}$ ,  $\beta_B$  by the method outlined in Section 3. The resulting matrices  $\Lambda_{mpA}$ ,  $\beta_A$  and  $\Lambda_{MPB}$ ,  $\beta_B$  are given in Table 25.
- D) Discard principal axes with low diagonal entries in the  $m{\beta}$  matrix. For study A discard the third column of the matrices  $\Lambda_{\mathrm{mpA}}$  and  $m{\beta}_{\mathrm{A}}$ . For Study B discard the fourth column of the matrices  $\Lambda_{\mathrm{MPB}}$  and  $m{\beta}_{\mathrm{B}}$ .
  - 1) Compute the square root of each remaining diagonal and record in the  $\sqrt{\beta}$  row.
  - 2) Compute the reciprocal of the  $\sqrt{B}$  and enter the result in the  $\sqrt[4]{B}$  row.
  - 3) Check the computations in 1 and 2 above by multiplying each  $1/\sqrt{\beta}$  by the corresponding diagonal entry. The product should approximate  $\sqrt{\beta}$  within  $\pm$  2 of the last deciral place carried.
  - 4) The matrices  $\beta_A^{\frac{1}{2}}$  and  $\beta_B^{\frac{1}{2}}$  are given in Table 26 and contain the  $1/\sqrt{\beta}$  as diagonal entries.
- E) Compute the matrix product  $T_{mpA} = A_{mpA}$   $\beta_A = for study A$ ,
  Table 27. (See Section 5, Paragraph c).
  - 1) Eliminate the column in  $\Lambda_{\rm mpA}$  corresponding to the principal axes discarded in D above. Since the third column of  $\beta_{\rm A}$  was eliminate the third column of  $\Lambda_{\rm mpA}$  will likewise be eliminated. The  $\Lambda_{\rm mpA}$  matrix will now have three rows and two columns.
  - 2) Multiply each entry in the first column of Ami-A by the

diagonal entry of the first column of  $\beta_A^{-\frac{1}{2}}$  and record in the corresponding cell of  $T_{mpA}$ . (Table 27).

- 3) Multiply the E entry in the first column of AmpA by the diagonal entry of the first column of  $\beta_A^{-\frac{1}{2}}$  and record in the Ch row of  $T_{mpA}^{-\frac{1}{2}}$ . Sum the entries in the first column of  $T_{mpA}^{-\frac{1}{2}}$  (exclusive of the Ch entry) and enter in E cell. This should agree with the Ch entry to within  $\pm 2$  of the last decimal place carried.
- -4) Compute the entries in the second column of  $T_{mpA}$  by using the entries in the second column of  $A_{mpA}$  and the diagonal entry of the second column of  $\beta_A^{-\frac{1}{2}}$ . Follow the procedure outlined in 2 and 3 above.
- F) Compute the matrix product  $T_{MPB} = A_{MPB} \beta_B^{-\frac{1}{2}}$  for Study B. (Table 27).

Follow procedure outlined in E above. The matrix  $\Lambda_{\text{MPB}}$  in the example has four rows and three columns after the fourth column has been eliminated.

- c. Compute transformations to congruent factors
  - A) Compute the matrix  $G = T_{mpA}^{*} (F_{JmA}^{*} F_{JMB}^{*}) T_{MPB}^{*}$  (See Table 28).
    - 1) Compute the matrix FinAFJMB
      - a) Compute the sum of products between the first column of  $F_{JmA}$  and first column of  $F_{JMB}$  (Table 23) and enter the result in the first cell of the first row of  $F_{JmA}^1F_{JMB}$
      - b) Compute the sum of products between each remaining column of  $F_{JmA}$  and the first column of  $F_{JmB}$ . Enter the result in the corresponding cell of the first column of  $F_{JmA}F_{JmB}$ .
      - c) Compute the sum of products between the entries in the  $\Sigma$  column of  $F_{JmA}$  and the first column of  $F_{JmB}$ . Enter result in the Ch cell of the first column of  $F_{JmA}^{\dagger}F_{JmB}$ .
      - d) Sum the entries in the first column of F'<sub>JmA</sub>F<sub>JMB</sub> (exclusive of the entry in the Ch cell) and enter the total in the Σ cell of the first column of F'<sub>JmA</sub>F<sub>JMB</sub>. This entry should agree with the Ch entry to within ±2 of the last decimal carried.

e) Second column of FinAFJMB

Compute the second column of  $F_{JMA}^*F_{JMB}$  by using the second column of  $F_{JMB}$  and repeating steps a-d.

f) Remaining column of Fina JMB

Compute the remaining columns of  $F_{JMA}^*F_{JMB}^*$  by using the corresponding columns of  $F_{JMR}^*$  and repeating steps a-d.

2) Compute the matrix product TimpA (FinAFIMB)

Using the columns of  $T_{mpA}$  and the columns of  $(F_{JmA}^*F_{JMB}^*)$  compute the matrix product according to the procedure outlined in the preceding step.

3) Compute the matrix product

$$G = T_{mpA}^{*} (F_{JmA}^{*} F_{JMB}^{*}) T_{MPB}$$
 (See Section 5 Paragraph c.)

a) Compute the sum of products between the entries in the first row of the matrix  $T_{mPA}^{i}$  ( $F_{JmA}^{i}F_{JMB}$ ) and the entries in the first column of  $T_{MPB}$  and record in the first cell of the first column of G. For example:

b) Compute the sum of products between the entries in the second row of  $T'_{MPB}$  ( $F'_{JmA}F_{JMB}$ ) and the first column of  $T'_{MPB}$ . Enter the result in the second cell of the first column of G. For example:

$$-.1941 = (-.2424)(.8287) + (-.1105)(.2814) + (.4702)(.2415) + (-.4207)(.1871)$$

- c) Compute the sum of products between the E row of T' mpA (F' mA JAB) by the first column of T mpB. Enter the result in the Ch cell of the first column of G. Sum the entries in the first column (exclusive of the Ch cell) and enter the total in the E cell. This entry should agree with the Ch entry to within +2 of the last decimal place carried.
- d) Compute the remaining columns of G by computing the sum of products between each row of  $T_{mpA}$  ( $F_{JmA}^{*}F_{JMB}$ ) and each remaining column of  $T_{MPB}$  and entering the result in the corresponding cell of G.

- B) If the number of principal axes p with significant latent roots  $\boldsymbol{\beta}_A$  for Study A is less than or equal to the number of principal axes P with significant latent roots  $\boldsymbol{\beta}_B$  for Study B, as is true for the example: (For detailed procedure see following paragraph D.)
  - 1) Compute H = GG
  - 2) Obtain latent vectors and roots of  $\Pi_A$ :  $\Lambda_{rA}$  and  $\mathcal{G}_r^2$  See Section 3.
  - 3) Discard vectors with low \$18.
  - 4) Compute

$$T_{\text{mrA}} = T_{\text{mpA}} \Lambda_{\text{rA}}$$

$$T_{\text{MrB}} = T_{\text{MPB}} G' \Lambda_{\text{rA}} g_{\text{r}}^{-1}$$

- C) If the number of principal axes p with significant latent roots  $\boldsymbol{\beta}_{A}$  for Study A is greater than the number of principal axes P with significant latent roots  $\boldsymbol{\beta}_{B}$  for Study B:
  - 1) Compute H<sub>B</sub> = G'G
  - 2) Obtain latent vectors and roots of  $H_B$ :  $\Lambda_{rB}$  and  $\mathscr{G}_r^2$
  - 3) Discard vectors with low pra.
  - 4) Compute

$$T_{\text{mrA}} = T_{\text{mpA}}G\Lambda_{\text{rB}} \phi^{-1}$$

$$T_{\text{MrB}} = T_{\text{MPB}} \Lambda_{\text{rB}}$$

- D) Since the number of principal axes with significant latent roots  $\beta_{\rm A}$  for Study A is less than the number of principal axes with significant latent roots  $\beta_{\rm B}$  for Study B, the equations in B will be solved as follows:
  - 1) Compute the mutrix product HA = GG \* (Table 29)
    - a) Compute sum of squares of first row of G and record in the first cell of the first row of  $H_{\underline{A}}$ .
    - b) Compute the sum of products between the first and second rows of G and enter in the second cell of the first column of  $\mathbf{H}_{\underline{z}}$

- c) Compute the sum of products between the first and E row of G and enter in the Ch cell of the first column of H<sub>A</sub>. To check the computations, obtain the sum of the first column of H<sub>A</sub> (exclusive of the Ch entry) and record in the E row. The Ch entry and E entry should agree within +2 of the last decimal place carried.
- d) Using the second row of G compute the second column of  $H_A$  as outlined in a to c above, the sum of squares of the second row being the diagonal of the second column of  $H_A$ .
- 2) Obtain the latent vectors and roots of  $H_A$  by the procedure outlined in Section 5. The resulting matrices are given in Table 29 as matrice  $A_{rA}$  and  $\beta_r^2$ .
- 3) Compute the square root of the diagonal entries in \$\phi\_r^2\$ and record in the √d row. Compute the reciprocal of √d and enter the result in the 1√d row. Check the computations by multiplying each 1√d by the corresponding diagonal entry. The product should approximate √d to within +2 of the last decimal place.
- 4) The entries in the 1/Vd row are the diagonal elements of the matrix  $\phi_{r}^{-1}$  given in Table 29.
- 5) Compute the matrix product  $T_{mrA} = T_{mpA} \Lambda_{rA}$  (Table 30)

  Using the rows of the matrix  $T_{mpA}$  and the columns of the matrix  $\Lambda_{rA}$ , follow computational procedure for multiplying matrices outlined in Section 5, paragraph c.
- 6) Compute the matrix product  $T_{MrB} = T_{MPB}G^{\dagger}\Lambda_{rA} \phi_{r}^{-1}$  (Table 31)
  - a) Compute the matrix product TurnG'
    - (1) Compute the sam of products between the first row of T<sub>MPB</sub> and first row of G and enter the results in the first cell of the first column of (T<sub>MPB</sub>G')
    - (2) Compute the sum of products between the remaining rows of T<sub>MPB</sub> and the first row of G and enter the result in the corresponding cell of the first column of (T<sub>MPB</sub>G')

- (3) Compute the sum of products between the  $\Sigma$  row of  $T_{MPB}$  and the first row of G and enter in the Ch row of the first column of  $(T_{MPB}G^*)$
- (4) Sum the entries in the first column (exclusive of the Ch cell) and enter the result in E cell of the column.

  This should agree with the Ch entry to within \*2 of the last decimal place carried.
- (5) Compute the remaining columns of (T<sub>MPB</sub>G') by using the rows of T<sub>MPB</sub> and each of the remaining rows of G and entering the results in the columns of (T<sub>MPB</sub>G') corresponding to the rows of G.
- b) Using the rows of the matrix  $(T_{MPB}G^*)$  and the columns of the matrix  $\Lambda_{rA}$  follow the computational procedure for multiplying matrices outlined in Section 5, Paragraph c, to obtain the matrix  $(T_{MPB}G^*)$   $\Lambda_{rA}$
- c) Using the rows of the matrix  $(T_{MPB}G^*)$   $\Lambda_{rA}$  computed in b, and the columns of the matrix  $\phi_r^{-1}$  (Table 29) follow the computational procedure for multiplying matrices cutlined in Section 5, Paragraph c to obtain the matrix

$$T_{\text{MrB}} = T_{\text{MPB}} G \cdot \Lambda_{\text{rA}} \phi_{\text{r}}^{-1}$$

- 7) Adjust columns of  $T_{
  m mrA}$  and  $T_{
  m MrB}$  to congruent factors:
  - a) For illustrative purposes the matrix T<sub>mrA</sub> and the matrix T<sub>MrB</sub> were copies in Table 32 from the results of previous computations (Tables 30, 31)
  - b) For each study compute the sum of squares for each column and enter the results in the row designated  $\Sigma \tau_{MTA}^{2}$  and  $\Sigma \tau_{MTB}^{2}$ .
  - c) For corresponding columns compute the average and enter the results in row designated  $\frac{1}{2}(\Sigma \tau_{\rm MrA}^2 + \Sigma \tau_{\rm MrB}^2)$ . For example the average between 1.356101 and 1.871665 is 1.613883
  - d) Compute the square root of the values obtained in (c)
  - e) Compute the reciprocal of the square roots computed in (d)

    Check the computations by multiplying each  $1/\sqrt{\frac{1}{2}(\Sigma \tau_{mvA}^2 + \Sigma \tau_{MrB}^2)}$ by the corresponding  $\frac{1}{2}(\Sigma \tau_{mvA}^2 + \Sigma \tau_{MrB}^2)$ . The product should

approximate  $\sqrt{\frac{1}{2}(E\tau_{mrA}^2 + E\tau_{MrB}^2)}$  to within \$2\$ of the last decimal place carried. These reciprocals of the square roots are the diagonal elements of the matrix  $D_r$  (shown in Table 52)

- f) Compute the matrix  $T_{mrA} = T_{mrA}D_r$  (Table 33)
- g) Compute the matrix T<sub>MrB</sub> = r<sub>MrB</sub>D<sub>r</sub> (Table 55)
- d. Compute loadings on congruent factors (Table 34)
  - A) Compute the matrix F<sub>JrA</sub> = F<sub>JrA</sub>T<sub>mrA</sub>

Using the rows of the matrix  $F_{JmA}$  (Table 25) and the columns of the matrix  $T_{mrA}$  (Table 33) follow the computational procedure outlined in Section 5, Paragraph c.

B) Compute the matrix  $\mathbf{F}_{JrB} = \mathbf{F}_{JMB}\mathbf{T}_{MrB}$ 

Using the rows of the matrix  $F_{\rm JMB}$  (Table 23) and the columns of  $T_{\rm MTB}$  (Table 33) follow the computational procedure outlined in Section 5, Paragraph c.

Table 21
REFERENCE FACTOR MATRICES
STUDY A

# P<sub>jmA</sub>

	1	11	111	Σ	47.
1	.35	34	.48	.49	4
2	.32	12	.34 .08	•54 •22	10
3	.54	40	.08*	.22	8
4	.54 .34	23	01	.10	4
5	.40	.12	26	.26	6
5 6	.27	30	-,28	31	4 6 8
v	•~ (		<b>_</b>		

### STUDY B

# F<sub>JMB</sub>

	I	II	III	IA	Σ	JB
1 2 3 1 5 6	.48 .48		.03 24 .29 .35 .36	.01 .06 12	.95 .40 1.04 1.16 .46	5 12 7 3 5 6

dy + dy	88888		
$a_{JB} = a_{JB} / \frac{1}{2} (a_{JA} + a_{JB})$	ાં ૧ ૧ & જે જે જે		•
ay = ay/ 2(ay + ayB)	.89 .91 41.1 41.1		•
2(oly + ols)	4.11.0 5.15.0 5.5.0 6.0	29.0	39.0
d JB	พนีะมพด		88
۹۶	കര്യകരമ		약
	•	៩	×

Table 25

### REFERENCE FACTOR MATRICES FOR THETS

### WITH ADJUSTED UNITS OF MEASURE

#### STUDY A

# P<sub>JMA</sub>

	1	11	111	CP.	Σ
1	.51	30	.43	.44.	.44
2	.29	11	.31	.49 .24	.49 .24
5	.58	45	•09	.24	.24
4	.39	26	01	.11	.12
2 3 4 5 6	.29 .58 .39 .44	.13	28	.26	.29
6	.31	34	32	35	-,35
Σ	2.32	-1.31	.22		1.25

### STUDY B

# r<sub>JMB</sub>

	I	11	III	14	CIP	Σ
1 2 3 4	.36	.36	.03	.51	1.05	1.06
2	.66	16	.03 26	.20	.44	.44
3	.46	.23	.27	.01	•97	•97 •99 •42
4	.41	.23	.30	.05	1.00	•99
5 6	.17	.03	.33 .36	11	.42	.42
6	.10	02	.36	25	.20	.19
Σ	2.16	.67	1.03	.21		4.07

STUDY A

y Jak 3-4

i ii iii Ch E

i .9584 -.5239 .0491 .4836 .4836
ii -.5239 .4871 -.1268 -.1636 -.1636
iii .0491 -.1268 .4700 .3923 .3923

### STUDY B

IV

# 

I	.9838	.2272	.1785	<b>.</b> 2250	1.6145	1.614
II	.2272	2623	.1862		.7708	
III	.1785	.1862	.4699	1513	.6833	.6833
TIF	9950	0061	_ 1513	2133	3801	.390

ш

I

11

### STUDY A

		Agga			•	A h	
	<b>p-1</b>	<b>p-2</b>	p-3	7	-1	p-2	p-3
1 11 111 E	.8275 5449 .1352 .4178	.2407 .1268 9023 5948	.5073 .8288 .2362 1.5723	p-2 .0 p-30		.0036 .4749 0007 .6891 1.4512	0021 0007 .1290

and the matter souls. The second distributions have a second and a second distribution of the second d

#### STUDY B

	A <sub>M2</sub>	P <b>B</b>			A B		
P-1	P-2	P-3	<b>P.</b> 4	P-I	P-2	<b>P-3</b>	P-4
I .8940 II .3036 III .2605 IV .2018 E 1.6599	2165 .2236 .8347 4542 .3876	.3674 7885 .0390 4917 8738	1374 4859 .4836 .7150	P-1 1.1639 P-2 .0041 P-3 .0030 P-4 .0028 _β 1.0788 1//β .9270	.0040 .5532 .0030 .0019 .7438 1.3444	.0050 .0030 .2085 .0001 .4566 2.1901	.0028 .0019 .0002 .0036

•	STUDY A $\beta_A^{-\frac{1}{2}}$	,		•		SURPLY B	
	<b>p-1</b>	p-2	•		P-1	P-2	P-5
p-1 p-2	.8751	1.4512		P-1 P-2 P-3	.9270	1.344	2.1901

Table 27

	<b>p-1</b>	p-2	Σ
1 11 111	.7225 4758 .1180	.3493 .1840 -1.3965	1.0718 2918 -1.2785
СР	.3648	8632	,.,
Σ	.3647	8632	4985

STUDY A

 $T_{mpA} = \Lambda_{mpA} \beta_A$ 

AND THE PROPERTY OF THE PROPER

I .8287 -.2911 .8046 1.3422
II .2814 .3006 -1.7269 -1.1449
III .2415 1.1222 .0854 1.4491
IV .1871 -.6106 -1.0769 -1.5004
Ch 1.5387 .5211 -1.9137
E 1.5387 .5211 -1.9138 .1460

A CONTRACTOR OF A STANK

# T<sub>Jm</sub>ÅT<sub>JIG</sub>

	I,	п	m	IA
1 11 111	.8355 4969 3171	.2955 2384 .1216	.4643 2540 2540	.0535 0616 .3065
Ch	.6557	.1785	0437	.2984
Σ	.6557	.1785	0457	.298

## Tmai(FjmiFjig)

	I	II	III	IA
p-1 p-2	.8775 2424	.5411 1105	.4263 .4702	.1041 4207
Ch.	.6351	.2306	.8965	3165
Σ	.6351	.2306	.8965	3166

# $G = T_{mpA}(F_{JmA}F_{JMB})T_{MPB}$

	P-1	P-2	P-3
p-1	.9456	.2619	.0413
p-2	1971	.8219	.4890
Ch	.7485	1.08 <del>38</del>	.5303
E	.7485	1.08 <del>3</del> 8	.5303

Table 29 HA - 00' **p-1** .9645 .0491 .0491 •9535 p-2 p-2 Ch E 1.0135 1.0136 1.0026 ε, p-1 p-2 .7470 .6648 1.0084 .0003 **√**₫ 1.0042 -9537 ¹//ā .9958 1.0485 B 1.0485

STUDY A

-.8873 -.8872 -.3014 Ch Σ

-.3014

Table 31

STUDY B

Table 31

Table 3

(T<sub>MPB</sub>G')(A<sub>rA</sub>)

I .5472 -.4991
II -.2297 -.6695
III 1.0021 .3351
IV -.7288 -.7775
Ch .5907 -1.6111
E .5908 -1.6110

 $T_{\text{MrB}} = (T_{\text{MrB}}G^*)(\Lambda_{\text{rA}})\phi^{-1}$ 

I .5449 -.5233
II -.2287 -.7020
III .9979 .3514
IV -.7257 -.8152
Ch .5883 -1.6891
E .5884 -1.6891

	. r <sub>est</sub>			T <sub>McB</sub>	
	. A B			A	В
i ii iii	.7719 2331 8402 3014	2194 .4538 -1.1216 8872	E IA III II	.549 2287 -9979 7257	9233 7020 .3514 8192 -1.6891
ET 2	1.356101	1.512057	ET HEB	1.871665	1.554680
$\frac{1}{2}(\Sigma  au_{ m mrA}^2$	2+ET <sub>MrB</sub> 2)	,		1.613683	1.533369
Va (E.T	A+DTNAB)		•	1.270387	1.238293
1/1/2(2	T 2+ET NrB	· ·		.787162	.807563
			ď		

Δ

A .7872

.8076

STUDY A

Tara -Tara<sup>D</sup>r

1 .6076 -.1772
11 -.1835 .3665
111 -.6614 -.9058
Ch -.2373 -.7165
E -.2373 -.7165

STUDY B

T<sub>MrB</sub> = T<sub>MrB</sub>D<sub>r</sub>

	٨	В
I	.4289	4226
II	1800	5669
III	.7855	.28 <b>38</b>
IA	5713	6584
Ch	.4632	-1.3641
Σ	.4631	-1.3641

OY B	atu		STUDY A  P <sub>JrA</sub> = P <sub>JmA</sub> T <sub>mrA</sub>		
Jus <sup>T</sup> al.	7 <sub>JrB</sub> = 1				
Į	A		3		
7 3 2 2 .0	0639 0066 .3623 .3415 .3896 .4721	1 2 3 4 5	5544 3725 3419 1553 .2233 .1103	0410 0086 .3718 .2913 .4287 .4624	123456
-1.1	1.4949		-1.0905	1.5045	Ch
-1.1	1.4950		-1.0905	1.5046	Σ

All the state of the state of

AGENERAL THE SECTION AND ASSESSED ASSESSED.

#### 2. Rotation of Axes in the Congruent Space

The illustration used in the preceding section will be used in this section also. One rotation of the axes in the congruent space will be described. As indicated in Section IV of the body of the report, the computations are of an identical nature each rotation from one set of axes in the congruent space to another set of axes. The forms given in this section can be used for all such rotations of axes. In the particular case illustrated and described, the rotation is from the congruent factors, r, to a rotated set, s. An identical computational procedure would be used in rotating from factors, s, to a set of factors,

Table 55 gives the transformation matrices for the desired rotation. It is assumed that the entries in the  $T_{rs}$  matrix have been derived from some procedure for deciding where to rotate the axes. In the present case these entries were determined graphically from a plot of the loadings on the two congruent factors in accordance with usual graphical rotation of axes procedures involving oblique axes. In the example, described in Section IV, the entries in the  $T_{rs}$  matrix were obtained from solving sets of simultaneous linear equations. Other methods might be employed to obtain the entrees in  $T_{rs}$ . The computational procedure to complete the rotation is described in the following steps.

a. Sum the columns of  $T_{rs}$ .

新されるなど。 このなかが変しなる対象

- b. Compute the matrix product  $T_{mrA}$   $T_{rs}$  by the procedure described in Section 5, Paragraph c. Record the results in matrix  $T_{msA}$ . Table 35 gives matrix  $T_{mrA}$  for the example. To obtain the checks, multiply the  $\Sigma$  row of  $T_{mrA}$  by the columns of  $T_{rs}$ . Sum the columns of  $T_{msA}$  (exclusive of the Ch. entries) and record in the  $\Sigma$  row. These sums should agree with the Ch entries within  $\frac{1}{2}$  in the last decimal place carried.
- c. Compute the matrix product  $T_{\rm MrB}\,T_{\rm rs}$  by the same procedure as in step b, recording the results in the matrix  $T_{\rm MsB}$ .

- d. Compute the sum of squares of the entries in each column of matrix  $T_{max}$  and record the results in the  $\Sigma T_{max}^2$  row.
- e. Compute the sum of squares of the entries in sech column of matrix  $T_{\rm MaB}$  and record the results in the  $\Sigma T_{\rm MaB}^2$  row.
- f. Sum each pair of corresponding entries in the  $\Sigma \tau_{\rm msA}^2$  and  $\Sigma \tau_{\rm NaB}^2$  rows, divide the sum by 2, and record in the  $\frac{1}{2}(\Sigma \tau_{\rm msA}^2 + \Sigma \tau_{\rm NaB}^2)$  row.
- g. Find the square root and reciprocal of the square root of each entry in the  $\frac{1}{2}(\Sigma \tau_{\text{msA}}^2 + \Sigma \tau_{\text{MsB}}^2)$  row and record the results in the  $\frac{1}{2}(\Sigma \tau_{\text{msA}}^2 + \Sigma \tau_{\text{MsB}}^2)$  and  $1/\sqrt{\frac{1}{2}(\Sigma \tau_{\text{msA}}^2 + \Sigma \tau_{\text{MsB}}^2)}$  rows. A check on these computations is to multiply each  $1/\sqrt{\frac{1}{2}(\Sigma \tau_{\text{msA}}^2 + \Sigma \tau_{\text{MsB}}^2)}$  entry by the corresponding  $\frac{1}{2}(\Sigma \tau_{\text{msA}}^2 + \Sigma \tau_{\text{MsB}}^2)$ , the product should equal the  $\sqrt{\frac{1}{2}(\Sigma \tau_{\text{msA}}^2 + \Sigma \tau_{\text{MsB}}^2)}$  within  $\frac{1}{2}$  in the last decimal place carried.
- h. Using the  $1/\sqrt{\frac{1}{2}(\Sigma \tau_{\rm maA}^2 + \Sigma \tau_{\rm MaB}^2)}$  for each column, multiply the entries in matrices  $T_{\rm rs}$ ,  $T_{\rm msA}$ ,  $T_{\rm MaB}$  in that column and record the results in the corresponding cells of the matrices  $T_{\rm rs}$ ,  $T_{\rm msA}$ , and  $T_{\rm MsB}$ . The Ch entry for each T matrix is obtained by multiplying  $\Sigma$  entry by the  $1/\sqrt{\frac{1}{2}(\Sigma \tau_{\rm maA}^2 + \Sigma \tau_{\rm MsB}^2)}$  for the column. Sum each column in each T matrix (exclusive of the Ch entry) and enter results in the  $\Sigma$  row. These entries in the  $\Sigma$  rows should agree with the corresponding Ch entries within  $\frac{1}{2}$  in the last decimal place carried.
- i. Obtain the sum of squares of entries in each column of the  $T_{msA}$  matrix and record in the  $\Sigma t^2$  row. Repeat for the  $T_{msB}$  matrix.
- j. Sum each pair of corresponding entries in the Et<sup>2</sup> and Et<sup>2</sup> rows; divide each sum by 2, and record in

- the  $\frac{1}{2}(Et_{msA}^2 + \frac{E^2}{N}MsB)$  row. These entries should equal unity within  $^{\frac{1}{2}}2$  in the last decimal place carried in the T matrices.
- k. Compute the matrices of loadings of the tests on the rotated factors,  $F_{JBA}$  and  $F_{JBB}$  (See Table 36). Obtain the matrix products  $F_{JmA}^{T}_{msA}$  and  $F_{JMB}^{T}_{MsB}$  by the procedure given in Section 5, Paragraph c. A computational short cut which is only slightly less accurate is to obtain the matrix products  $F_{mrA}^{T}_{rs}$  and  $F_{MrB}^{T}_{rs}$ .

	$ au_{rs}$			Tr	•
		b			b
A B	1.00 06	.37 -1.00	A B	1.0276 0617	.4169 -1.1267
			Ch	.96 <del>59</del>	7098
Σ	.94	63	Σ	.9659	7098
	T <sub>msA</sub> = T <sub>m</sub>	ra Tra		T <sub>me</sub>	
		ъ		<b>a</b>	ъ
1 11 111	.6182 2055 6071	.4020 4344 .6611	1 11 111	.6352 2112 6238	.4529 4894 .7449
Ch	1943	.6287	Ch	1998	. 7084
Σ	1944	.6287	Σ	1998	.7084
Et <sup>2</sup> m MaaA	.7929 <b>72</b>	.787361	Et <sup>2</sup> m msA	.837211	.999 <b>507</b>
	T <sub>MsB</sub> -	T <sub>M B</sub> T <sub>xs</sub>		T <sub>M</sub>	a <b>B</b>
	a	ď		a	ъ
IV III I	.4543 1460 .7685 5318	.5813 .5003 .0058 .4470	IA III I	.4568 1569 .7897 5465	.6550 .5637 .0077 .5036
Ch	.450	1.5354	Ch	.5600	1.7500
Σ	.5450	1.5354	Σ	.5600	1.7300
er M <sup>e</sup> MoB	1.101108	.783065	H Et <sup>2</sup> K	1.162691	1.000455
$\frac{1}{2} ( E r_{\text{meA}}^2 + E r_{\text{MaB}}^2 )$	.947040	.787715	T (Et 2 HEBA HE L 2)	.999951	.999981
1 / (27 A+Z+Z) m mgA M MaB)	.973160	.887532			
1 $\sqrt{\frac{1}{2}(2\tau^2_{\text{meA}} + 2\tau^2_{\text{M}})}$	1.027580	1.126720			

STUDY A

----

Tak - Tak Tank

	•	ъ
1 2 3 4 % 6	0080 .0141 .4031 .3089 .4267	.6075 .4161 .5402 .2964 0729
6	.4683	.0684
Ch	1.6131	1.8557
Σ	1.6131	1.8557

STUDY B

F<sub>JuB</sub> = F<sub>JMP</sub>T<sub>KuB</sub>

b

1 -.0317 .5951
2 .0175 .4408
3 .3680 .4381
4 .3665 .4257
5 .3956 .0754
6 .4706 -.0689
Ch 1.6064 1.9062
E 1.6065 1.9062

#### 5. Determination of Non-Congruent Axes

The congruent factors established by the preceding procedures will usually be fewer than the total number of factors in either study. A set of non-congruent axes is to be established in each study. The number of congruent factors and non-congruent axes is to equal the total number of factors in the study. In Study B of the pair of studies used in the body of this report there was a total of 12 factors. A set of 5 congruent factors were determined. This left 7 factors to be established as non-congruent axes. The computing procedure for establishing the non-congruent factors follows. These directions will be illustrated by a fictitious example for which there is a total of rive factors. Table 37 gives the transformation T<sub>mrA</sub> to three congruent factors. Two non-congruent factors are to be established.

- a. Prepare a work sheet like the one given in Table 58.
  - 1) Sections A, B, and C are to have a row for each reference factor.
  - 2) Sections A and D are to be located vertically from each other and are to have as many columns as there are factors in the study. Head the columns 1, 2, 3, etc.
  - 3) Sections B and K are to be located vertically from each other and are to have as many columns as there are factors. Head the columns 1', 2', 5', etc.
  - 4) Section C is to have columns for the non-congruent axes. Head the columns 4", 5", etc. The first number in this series is one greater than the number of congruent factors. The last number in the series is the number of reference factors in the study.
  - 5) Sections D and E are to be located horizontally from each other and are to have as many rows as there are factors in the study. Head the rows 1', 2', 3', etc.
  - 6) Record unities in the diagonal cells, from upper left to lower right, of Section D. Make dashes in all other cells of Section D.
  - 7) Make deches in the diagonal cells, from upper left to lower right, and in all cells to the right of the diagonal of Section E.
- b. Copy the matrix  $T_{mrA}$  (or  $T_{MrB}$ ) into the left portion of Section A using as many columns as there are congruent factors. The remaining columns will be left blank for the present. Enter previously determined column totals of  $T_{mrA}$  in the  $\Sigma$  row of Section A. Check the copying by summing the columns of Section A, these sums should agree with the previous totals entered in

the E row.

- c. Obtain column 1' of Section B.
  - 1) Copy column 1 of Section A into column 1' of Section B. Enter the sum of column 1 from the E cell of column 1 into the Ch cell of column 1'. Sum the column 1'(exclusive of the Ch entry) and record th, result in the E cell of column 1'. The entries in the Ch cell and E cell of column 1' should agree.
  - 2) Find the largest number, ignoring sign, in column 1' of Section B (exclusive of the Ch and E row entries). Underline this number and make dashes in the remaining cells of the row of Section B in which the number is located. This number is the 1' pivot entry. The row is the 2' pivot row. In the example the largest number in column 1' of Section B is .6612 in row iv. This is the 1' pivot entry and row iv is the ' pivot row.
- d. Compute column 2' of section B:
  - 1) Compute the entry in Section E row 2' column 1' by dividing the entry of Section A column 2 in the 1' pivot row by the 1' pivot entry and recording the result with reverse sign in the Section E row 2' column 1' cell. In the example, the entry of Section A column 2 in the 1' pivot row is .3421. Division of .3421 by .6612, the 1' pivot entry, yields .5174 which is recorded in Section E row 2', column 1' with a negative sign.
  - 2) Compute each entry in Section B column 2' by multiplying each row of Sections A and B by Sections D and E row 2'. For the entry in Section B row i column 2', multiply rowi of Section A and B by row 2' of Section D and E. Only columns 2 of Section A and 1' of Section B will be involved since row 2' of Sections D and E has entries in these columns. In the example, row i

(.4175)(1.0000) + (.1393)(-.5174) = .3454.

Note that the entry in the 1' pivot row should be zero and need not be recorded. If this entry in the 1' pivot row is not zero, the entry in Section E row 2' column 1' is incorrect.

5) Compute product of Σ row of Sections A and B and row 2' of Sections D and E and record the result in the Ch cell of Section B column 2'.

- 4) Sum the entries in Section B column 2' (exclusive of the Ch entry) and record the result in the E row. The Ch and E entries should agree within \$2 of the last decimal place carried.
- 5) Select the largest entry, ignoring sign, in Section B column 2' (exclusive of Ch and E entries). This is the 2' pivot entry and is to be underlined. Make dashes in the remaining cells of Section B in the row to the right of the 2' pivot entry. The row containing the 2' pivot entry is the 2' pivot row.
- s. Compute column 3: of Section B:
  - 1) Compute the entry in Section E row 5' column 1' by dividing the entry of Section A column 3 in the 1' pivot row by the 1' pivot entry and recording the result with opposite sign in the Section E row 3' column 1' cell. In the example, -.7365 is the entry of Section A column 3 in the 1' pivot row (row iv which contains the underlined 1' pivot entry in Section B column 1'). Then:
    - -(-.7365)/.5612 = 1.1139.
  - 2) Multiply the 2' pivot row in Sections A and B by the portion of Sections D and E row 3' that has been determined, divide the result by the 2' pivot entry and record the result with reverse sign in the Section E row 3' column 2' cell. For the example, multiplication of Sections A and B row iii (the 2' pivot row) by Sections C and D row 3', division by the 2' pivot entry, and reversal of sign yields:
    - [(.0349)(1.6000) + (-.5023)(1.1159)] / .9481 = (-.5246)/ .9481 = .5535.

This result is recorded in Section E row 3' column 2'.

5) Multiply each row of Sections A and B by row 3' of Sections D and E and record the results in Section B column 3'. In the example, multiplication of row i of Sections A and B by row 3' of Sections D and E yields;

(.3571)(1.0000) + (.1393)(1.1139) + (.3454)(.5533) = .7034 which is recorded in Section B row 1 column 3'. Note that the entries in the 1' and 2' pivot rows should be zero. If either of these entries is not zero, the entries in Section E row 3' are incorrect.

- 4) Multiply the E row of Sections A and B by the 5' row of Sections
  D and E and record the result in the Ch cell of Section B column 5'.
- 5) Sum the entries in Section B column 3' (exclusive of the Ch row) and record the result in the E row. The Ch and Σ entries should agree within ±2 of the last decimal place carried.
- 6) Select the largest entry, ignoring sign, in Section B column 3'
  (exclusive of Ch and E entries). This is the 3' pivot entry and
  is to be underlined. Make dashes in the remaining cells of Section
  B in the row to the right of the 3' pivot entry. The row containing
  the 3' pivot entry is the 5' pivot row.
- f. Compute remaining columns of Section B corresponding to columns of Section A containing congruent factors. In the present example there were three congruent factors and the computations of columns of Section B will stop, therefore, with column 3'. If the matrix T<sub>mrA</sub> had seven columns, the computations would continue through column 7' of Section B. (During the process of determining non-congruent axes, corresponding columns of Section B will be determined. In the example the non-congruent factor, later recorded in column 4 of Section A, was used in obtaining column 4' of Section B. These steps are subsequent to the present step.) Follow the procedure outlined in the foregoing step e. For each additional column added to Section B there is an additional entry in the corresponding row of Section E. Step e3 gives the general procedure for determining the entries in the Section E row.
- g. Determine the first non-congruent axis.
  - 1) Record unity in the first column of Section C in some row that is not a pivot row. In the example column 4: had not been recorded in Section B and row i was not a pivot row. Unity was recorded in row i of Section B column 4". (The row selected is likely to become the next pivot row when computations return to Section B.)
  - 2) Record zeros in the first column of Section C in other rows that are not pivot rows. In the example, .0000 was recorded in row ii of Section C column 4".
  - 5) Multiply the first column of Section C(using those entries already recorded) by the last column recorded in Section B, divide the resulting sum of products by the pivot entry in the column of Section B, and record with reverse sign in the corresponding pivot row in

Section C first column. In the example:

- a) Section C column 4" was multiplied by Section B column 5', yielding:
  - (.7054)(1.0000) + (.4984)(.000) = .7054.
- b) This result was divided by the 5' pivot entry:

#### .7034/(-1.0263) = -.6854

- c) The sign of the result was changed and the .6854 was recorded in Section C column 4" in row v. the 5' pivot row.
- 4) Multiply the first column of Section C by the next to last column recorded in Section B, divide the resulting sum of products by the pivot entry in the column of Section B, and record the result with opposite eign in the pivot row in the first column of Section C. In the example, column 4° of Section C was multiplied by column 2° of Section B.

The result was divided by the 2'pivot entry (in rew lii) and this result was recorded with emposite sign in Section C column 4" row ili, the 2' pivot row.

NAMES AND REPORTED REPORTED (BETTERSTORY FOR SOME RESERVED BARROOM, RECTORS

- -[(.3454)(1.0000) + (-.6138)(.0000) + (-.4114)(.6854)] /.9481
  - = (.0631.2644)/.9481 = -.0669
- 5) Continue the process described in steps 3 and 4 working back one column of Section B each time and recording the result in the pivot rew in the first column of Section C. All entries in the column of Section C will then be determined.
- 6) Sum the first column in Section C and record the result in the E row.
- 7) Obtain the sum of squares of the entries in the first column of Section C (exclusive of the E entry) and record the result in the

Era row.

- 8) Obtain the square root of the entry in the  $\mathbb{E}e^2$  row, recording in the  $\sqrt{2}e^2$  row, and the reciprocal of the square root, recording in the  $1/\sqrt{2}e^2$  row. These computations may be checked by multiplying the  $1/\sqrt{2}e^2$  entry by the  $\mathbb{Z}e^2$  entry, the result should agree with the  $\sqrt{2}e^2$  entry within  $\pm 2$  of the last decimal place carried.
- 9) Multiply each entry in the column of Section C by the lyvic entry and record in the convergenting column of Section A. Pecults from cell in '" of Section C were recorded in column 4 of Section A. To check these entry in the column of Section C

by the  $1\sqrt{ID^2}$  and record in the Ch row of the Section A column. Sum the entries in the Section A column (exclusive of the Ch entry) and record in the E row. The Ch and E entries should agree within  $\pm 2$  of the last decimal place carried.

- 10) Obtain the sum of squares of the entries in the new Section A column and record the result in the Ets row. This entry should equal unity within +1 in the last decimal place carried.
- 11) Check that the product between the new column of Section A and and each preceeding column of Section A is zero within 11 of the last decimal place carried.
- h. Compute a new column of Section B for the column added to Section A.

  Follow the procedure outlined in step e. Also see comments on procedure
  in Section f.
- i. Compute a new column of Section C for the next non-congruent exis. Repeat step g interpreting the directions to indicate the second non-congruent axis werever the first non-congruent exis is mentioned.
- j. Repeat steps h and i for subsequent non-congruent axes until all columns of Section A are completed. The columns added to Section A contain the direction cosines of the non-congruent axes. These columns may be copied into a  $A_{mid}$  (or  $A_{Min}$ ) matrix such as is given in Table 6.

## TRANSFORMATION TO CONGRUENT PACTORS FOR A PICTITIOUS EXAMPLE

	TarA				
Reference	Congruent Factors				
Factors			<del></del>		
1	.1393	.4175	.3571		
11	.4871	<b>3</b> 618	.2954		

-.502**3** .6612 111 -.5218 -.2153

-.7365 -.5611 -.6102 .5642 .5720 Σ

	Section C		1.000 1.000 1.000 1.000 1.000 1.000 1.000	2.1.2 1.98.1.1 2.05.7.		
	Bect	1	1.000 .000 .000 .000 .000 	1.5781 2.2.1.4.759 1.512149 1.512149 1.518149		
Table 38	Section B	1, 2, 3; 4; 5°	1393 .3454 .7034 1.2148 4671 .6138 .4994 .2243 .5023 .9-81 .6612 .4114 -1.0263	.5720 .2682 .1754. 1.4592 .5720 .2683 .1755 1.4591	Section B	.517k 1.1139 .5533 .0504 .0348 .5053
	Bectlo. A	$(T_{merA})$ $(A_{minA})$	23 . 1775 . 3571 . 88 711 - 3518 . 2574 . 05 23 . 56.2 . 0519 - 05 24 - 5521 - 1525 - 53 25 - 5525 - 5525 - 5525	1.2589 1.5203 .5720 .56426102 1.2989 1.5303 1.0000 1.0001	Section D	1,000
			កជុដ្ឋា	世 克		びれながに

-101-

..1846 1.0000 8705 0793 .5591

2.1245

1.288. 1.388. 204.

MESSESSE REPORT NO SOUND DESCRIPTIONS RECEIVED NO CONTRACT DE PROPERT

#### 4. Determination of Latent Roots and Vectors

Given a symmetric matrix A<sub>O</sub>, such as in Table 39, it is desired to compute a matrix of latent vectors and the corresponding latent roots. The matrix A<sub>O</sub> in Table 43 in a close approximation of the matrix of latent vectors for the example. The diagonal entries in matrix A<sub>O</sub> of Table 43 are close approximations to the latent roots for the example. The method to be described is an adaptation of the successive rotations method developed by Trunan L. Kelley (1<sup>th</sup>). Each rotation transforms the axes to closer approximations to the latent vectors, with any desired degree of precision being obtained by taking more rotations. It may be necessary to carry more decimal places in later rotations to realize the potential precision. In the example only two rotations were computed to obtain a fair degree of precision. Most salient features of the procedure are illustrated however.

Tables 40 and 41 contain the computations for the first rotation and Tables 42 and 45 contain the computations for the second rotation. Each rotation starts from an A matrix and produces a revised A matrix. Botation 1 started from the An matrix of Table 39 and produced the A matrix in the lower right of Table 41. Rotation 2 started from the A, matrix and produced the A, matrix in the lower right of Table 45. Note that each of these A mitrices is symmetric (each row of the untrix has idnetical entries with those of the corresponding column). The largest entries are in the diagonal from upper left to lower right and from rotation to relation the off-diagonal entries are becoming smaller. The solution occurs when the off-diagonal entries become zero. The diagonal entries are then the latent roots. Each rotation has a mairix relating the A setrix produced by that rotation to the original A matrix. The A of matrix for the first rotation is in the upper left of Table 41. The A OR matrix for rotation 2 is in the upper left of Table 45. When the off-diagonal entries in the A matrix are zero, the A matrix contains the latent vectors.

In the following directions, only one rotation will be covered explicitly. It is expected that as many such rotations will be taken as naceseary to obtain the precision desiral for any particular solution.

- a. For each rotation prepare a set of work sheets. Tables 40 and 42 illustrate the set up for Work Sheet 1, Tables 41 and 43 illustrate the set up for Work Sheet 2. Each matrix in these work sheets is to be the same size as the matrix A<sub>n</sub>.
- b. Obtain the entries in the B and C matrices from the preceeding A matrix. For Rotation 1 the  $A_0$  matrix is used, for Rotation 2 the  $A_1$  matrix is used, etc.
  - 1) Find the largest off-diagonal entry irrespective of sign below the diagonal of the preceding A matrix. In Rotation 1 of the example, the largest entry, ignoring sign, below the diagonal of matrix  $\lambda_0$  in Table 39 is .58 in row 5 and column 2.
    - a) Subtract the diagonal entry for the row of the selected off-diagonal entry from the diagonal for the column of the selected off-diagonal entry. For the example: the diagonal of row 5 is 1.55, the diagonal of column 2 is 1.21, subtracting 1.55 from 1.21 yields -.34.
    - b) Divide the difference in diagonal entries by the off-diagonal entry, ignoring the sign of the off-diagonal entry. Record the result in the cell of matrix B corresponding to the selected off-diagonal entry. Note that the sign of the result will depend only on the cign of the difference between the diagonal entries; thus, if the diagonal entry nearer the upper left is larger than the diagonal entry toward the lower right, the sign is plus; if the reverse is true, that is of the two diagonal entries, the second diagonal entry from upper left to lower right is the larger, the sign is negative. For the example:

which is recorded in the matrix  $B_{01}$  ros 5 and column 2 cell. The sign is minus because the second diagonal entry of matrix  $A_0$  is less than the fifth diagonal entry. The entry in row 3 and column 1 of matrix  $B_{01}$  illustrates the case when the entry is positive. The first diagonal entry of matrix  $A_0$  is larger than the third diagonal entry of matrix  $A_0$ . The sign of the -.10 in row 3 and column 1 of matrix  $A_0$  is ignored.

From Table 44 find the c value corresponding to the b entry in the B matrix. The sign of the entry in the B matrix is ignored in finding the corresponding c value. For the example: the .59 from row 5 and column 2 of matrix  $B_{01}$  is in the interval of b values of Table 44 of .57 to .59 for which the corresponding c value is .75. In case the entry in the B matrix is plus, record the c value found in step c in the corresponding off-diagonal cell of the matrix C with the sign of the corresponding entry in the A matrix. Record unities in the corresponding diagonal calls of the C matrix and cory the off-diagonal entry reversing sign, into the symmetrical cell above the disgunal. For an example note the entries in rows 1 and 3 and columns 1 and 3 of matrix Co, in Table 40. In case the entry in the B matrix is minus, record the c value in the two corresponding diagonal cells of the matrix C. Record unity in the corresponding off-diagonal cell of matrix C and assign it the same sign as the off-diagonal entry in matrix A, and record unity with the opposite sign in the symmetric cell above the diagonal of matrix C. This step was followed for the

If the off-diagonal entry in a C matrix is .30 or larger, ignoring sign, the rows and columns of the entry are to be excluded from further consideration in determining other entries in the B and C matrices. This is true for the example where the largest off-diagonal entry in matrix  $\mathbf{A}_{\Omega}$ was in row 5 and column 2 which yielded off-diagonal entries of unity in matrix Co1. The rows and columns 2 and 5 were then excluded from the following steps in determining other entries in matrix Col. In the case of the second rotation none of the off-diagonal entries in matrix C10 of Table 42 were .30 or larger and no rows or columns were excluded from further consideration in determining other entries in the C, matrix. As a consequence, several off-disgonal entries appear in each column of matrix C12. Whenever the C matrix of f-diagonal entry is .30 or larger, only this off-diagonal entry should appear in its row or column of matrix C.

entries in rows 2 and 5 and columns 2 and 5 of the matrix Co1

in Table 40.

2) Select the next to the highest off-diagonal entry, irrespective of sign, below the diagonal of the A matrix. Do not consider any entry in a row or column that has been excluded from further consideration in step 1-f. In the example, in Rotation 1 rows 2 and 5 and columns 2 and 5 were excluded from further consideration in step 1-f. The next highest off-diagonal, irrespective of sign, in matrix A<sub>0</sub> of Table 39 in the rows and columns 1,3, and 4 remaining is -.10 in row 3 and column 1. The -.11 in row 5 and column 3 is not to be considered. When the off-diagonal entry has been selected, follow the procedure of steps 1-a to 1-f in obtaining the entry in matrix C.

In Rotation 2, the highest off-diagonal entry irrespective of sign of matrix  $A_1$  of Table 41 was -.10 in row 4 column 1. This yielded an off-diagonal matrix  $C_{12}$ , of Table 32, entry of -.12. Since this entry was not large enough, ignoring sign, to cause the exclusion of rows and columns 1 and 4, the entry of .09 in row 4 and column 2 of matrix  $A_1$  could be selected second. This yielded an entry of .11 in row 4 and column 2 of matrix  $C_{12}$ . The selection of entries in the A matrix, from high to low, and subsequent determination of entries in the C matrix is to continue until there are no more off-diagonal entries in matrix A from which to select.

A special case exists when a C matrix entry of .30 or greater, irrespective of sign, occurs on later selection of off-diagonal entries in matrix A and there is already an entry in the row or column of this entry in the C matrix. In this case, this entry is not to be recorded and the rows and columns in which it is located are to be excluded from further consideration in the selection of siditional entries in matrix A.

c. Compute the entries in the T matrix. In Rotation 1, Table 40, matrix T<sub>O1</sub> is identical with matrix C<sub>O1</sub> which is, therefore, to be copied. In all subsequent rotations the T matrix is to be obtained by multiplication of of the preceding Amatrix by the present C'matrix (see Section 5, paragraph c for the procedure in multiplication of metrices). In Rotation 2, A<sub>O1</sub> for Rotation 1 in Table 41 is multiplied by matrix C<sub>12</sub> in Table 42 to produce matrix T<sub>O2</sub> of Table 42. In obtaining the Ch row of matrix T, multiply the E row of the preceding A matrix by the columns of the C matrix.

For Rotation 1, the sums of the columns of the C matrix are to be recorded in the Ch row of matrix T<sub>O1</sub>.) The sums of the columns of the T matrix (exclusive of the Ch entries) are to be recorded in the E row and should agree with the entries in the Ch row within ±2 of the last decimal place carried. After the matrix T has been determined, the rows are to be summed and the results entered in the E column. The sums can be checked by summing both the E row and the E column. These two sums should agree.

Compute the E matrix. For an example see Table 42.

- 1) Obtain the first column of the E matrix.
  - a) Obtain the sum of squares of the entries in the first column of matrix T and record the result in the first row and first column of matrix E.
  - b) Obtain the sum of products between the entries in the first column and second column of matrix T and record the result in the second row and first column of matrix E.
  - c) Obtain the sum of products between the entries in the first column and each of the remaining columns of Matrix T and record the results in the corresponding rows and column 1 of matrix E.
  - d) Obtain the sum of products between the entries in the first column and the Σ column of matrix T and record the result in the Ch row and column 1 of matrix B.
  - a) Sum the entries(exclusive of the Ch entry) in the first column of matrix E and record the result in the Σ row and column 1. The Ch entry and the Σ entry should agree within ±2 in the last decimal place carried.
  - f) When the entries in the first column of matrix E have been checked in step e. these entries may be copied into the first row of matrix E.
- 2) Obtain each of the other columns of the matrix E by finding the sums of products between entries in the corresponding column and each of the other columns of matrix T. The sum of squares of entries in each column of matrix T is to be recorded in the corresponding diagonal cell of matrix E.
- 3) Pouble each diagonal entry in matrix Z and record the result in the corresponding cell of column 20; To check this work obtain the sum of the diagonal entries, double this sum and record in the Ch cell of

column  $2e_{jj}$ . Then sum the entries in the column  $2e_{jj}$  and record the result in the  $\Sigma$  cell of the column. The entries in the Ch cell and  $\Sigma$  cell should agree.

- 4) Obtain the square root of each diagonal entry in matrix E and record the results in the  $\sqrt{e_{kk}}$  row.
- 5) Obtain the reciprocals of the square roots determined in step 4 and record in the  $1/\sqrt{e_{kk}}$  row. Multiply each diagonal entry by the corresponding  $1/\sqrt{e_{kk}}$ . This product should agree with the  $\sqrt{e_{kk}}$  within  $\pm 2$  in the last decimal place carried.
- e. Obtain the F matrix. For an example see Table 42.
  - 1) Copy the  $1/\sqrt{e_{kk}}$  entries into the diagonal cells of matrix F.
  - 2) Multiply each off-diagonal entry in matrix E by the 1/√e<sub>kk</sub> at the bottom of the column containing the entry, divide this product by the 2e<sub>jj</sub> at the right of the row containing the entry, and record the result with reverse sign in the corresponding cell of matrix F.
    For example, the entry in row 2 and column 1 of the matrix E<sub>2</sub> in Table 42 is multiplied by the 1/√e<sub>kk</sub> entry at the bottom of column 1:

-.0110 x .9726 = -.01069860, this result is divided by the  $2e_{11}$  in row 2:

-.0106980/2.2140 = -.0048

which is recorded with opposite sign in row 2 column 1 of the matrix  $\mathbf{F}_{p}$ .

- 3) Obtain each entry in the Ch column by:
  - c) Multiplying by 3 the square root of the diagonal entry in the corresponding row of matrix E (this square root can be found in the  $\sqrt{e_{kk}}$  row in the column of the diagonal entry),
  - b) Subtracting from the product of step a, the sum of products between the entries in the row of matrix E and the entries in in the  $1/\sqrt{e_{kk}}$  row (including the diagonal entry), and
  - c) Dividing this result by the Rejj for the row.

    The computations for the first Ch entry in matrix F<sub>2</sub> in Table
    42 are derived from the first row of matrix E<sub>2</sub>.

## 2.04966926/2.1142 = .9695

(The subtractions in step b may be accomplished by adding the products with the signs of the entries in the F matrix row reversed.)

- Sum the entries (exclusive of the Ch entry) in each row of the matrix F and record the result in the  $\Sigma$  column. The Ch entry and the  $\Sigma$  entry should agree within  ${}^{\frac{1}{2}}$ 2 in the last decimal place carried.
- Compute the A matrix by multiplying the T matrix by the F matrix. (See Section 5, paragraph c for the procedure in multiplication of matrices.) In the example matrices T<sub>O2</sub> and F<sub>2</sub> are in Table 42. Their product is matrix A<sub>O2</sub> of Table 45. The columns of the A matrix are checked by multiplying the Σ row matrix T by matrix F and recording the results in the Ch row of the A matrix. The entries (exclusive of the Ch row entries) in the columns of the A matrix are to be examed and the results entered in the Σ row. The Ch entries and the Σ entries should agree within ½ in the last decimal place carried. Sum the rows of the A matrix and enter the results in the Σ column.
- g., Compute the entries in the A'A'matrix following the procedure given in step d for the E matrix. The diagonal entries in the A'A matrix should be unity within \$2\$ in the last decimal place carried in the Amatrix and the off-diagonal entries should be zero with this same degree of accuracy.
- h. Compute the matrix product A<sub>O</sub> A by multiplying the A<sub>O</sub> matrix by the A matrix. (See Section 5, paragraph c for the procedure in multiplication of matrices.) Check the rows of the product matrix A<sub>O</sub> A by multiplying each row of the matrix A<sub>O</sub> by the E column of the Amatrix and recording the results in the Ch column. Sum the entries (exclusive of the Ch entry) in each row of the product matrix A<sub>O</sub> A and record the result in the E column. The Ch and E entries should agree within \*\*2\* of the last decimal place carried.
- 1. Compute the new A matrix.
  - 1) Multiply each column of the product matrix A<sub>O</sub>A by the first column in the A matrix and record the result in the cell of the first row of A corresponding to the column of A<sub>O</sub>. Also multiply the Σ column of A<sub>O</sub> by the first column of the Amatrix and record the result in the first row, Ch column of matrix A. Sum the entries (exclusive of the Ch entry) in the first row of the new A matrix and record the result in the Σ column. The Ch entry and the Σ entry should agree to within †2 of the last decimal place carried.

2) Compute the remaining rows of the new A matrix using the corresponding columns of the A matrix and following the procedure described in step 1. The A matrix should be symmetric if enough decimal places were carried in the product matrix A<sub>O</sub> A. It would be preferrable to carry sufficient decimal places in this matrix product to ensure that the matrix A would be symmetric. Then, when one row of matrix A has been computed and checked, the entries in that row may be copied into the corresponding column.

# Table 59

AN ILLUSTRATIVE SYMMETRICAL MATRIX

			O		
	1	2	3	<b>h</b>	5
•	1.99 02 10 08 .00	02 1.21 .01 .02 .58	10 .01 1.69 .08 11	08 .02 .08 1.20	.00 .58 11 .10 1.55

					-112-					
			ы	2.3445 .5717 .9313 1.2896 2.0803					-	
			€	2,3445 .5717 .9313 1,2896 2,0803			ы	2.3 86.1 86.1 87.		
			7	0160 0470 0440 0440 0440			g	4.0.4.4.6.5.4.6.5.4.6.5.4.4.4.6.5.4.4.4.6.5.4.4.4.4		
			. <del>4</del>	0800 .0800 .12000		0 A <sub>01.</sub>		4864£		
		<sup>λ</sup> ο <sup>Δ</sup> 01	Ao Aoı	ю	.48110038 1.5934 .0536 1.		A1 = A'01 A0	ਰ ਰ	5.8.8.4	
					•	ه ۱	4 4 8	88.1.		
			Q	0120 1.1900 0820 .0920 1.5830		⋖ ′		48888		
	TIOM 1			н	1.9394 0221 5861 1000 0319			н	2.00 100 100 100	
Table 41	WORK SHEET 2, ROTATION 1						ገመከታሪ		•	
A			អ	1.25 20 1.00 1.40			2	.0000 .0000 .0000 .0000	0000	
	•		7	8. 8.	20					
			<b>#</b>	8.	98		4	0000. 0000. 0000.	1,0000	
		Tol Fl	ю	83 88 1.	1.25 l. 1.25 l.	το <sub>γ</sub> .	11)	.0000 1.0057 .0000	1.0057	
		Δ ol - Tol	Q	8 8	1 04.1	A OL	Ø	.0000	1.0000	
			н	8. 63.	.67		н	1.0057	1.0057	
	•			このされら	គូผ			икары	គូង	

ANTON BURNELLANDE

And the same of the

8.88.89. 8.88.89.89.

%%% %%% %%% %

8.888888 8.88888

988. 98. 98. 98. 98.

9621 9621 0155 0008

9499. 9799. 9899. 9899.

	a
Žį.	HOTATION
je je	ij
Table	SHEET
	HOPE

	H	Table 42						
	WORK SHEE	WORK SHEET 1, ROTATION 2						
				*40	2 - 10 TO	.8		
v			H		'n	.at	₩.	20.33
		40	1.0571	0110	2030	0203	.0080	2.1142
		ı ıo.	0303	898	1.0803	0357	958	2.1606
;		4 W	.0203	818	0337	1.0465 0080	1.0146	% % % % %
		g ;	1.0641	1.0882	1.0777	.9917	888	10.6110
<b>~</b>			1.0232	1.0%	1.039	1.0230	1.0073	or no or
.83		1/4-14	.9726	.9505	18%	<i>STT8</i> :	.988	
ខ្ទុខ					₽ <sub>CV</sub>	Ļ		
8.6				ŭ	$(x_{tk} - 1/N_{tk})$	(#     		
				· (Fage	( x 3 x - 3 x / 2 e 3 3 V Ex	.: 4		3 4 K)
ľ	н		п	લ	n	4	<b>.</b>	đ

वृह्यसम् ឧଞ୍ଛଞ୍ଛଞ୍ ଖ୍ୟ 84488 EE B12 (bg. - 7tx - 313) **७७**४५३ सम v 84849. લ ઇંડ્રિયું મુંઠ <sup>મુ</sup> લ श्रुश्चंत्रश्च इ.स 3.88 ទង់ខង់ន %ខ្លួំដូដ ಇំ។

WORK SEEM 2, ROTATION 2

		-114-		
	ы	1.9687 .8336 .5042 1.3236 2.5388	н	2.0153 2.0155 11.6649 1.2209 1795
Ao 4 oz	g	1.9687 .8336 .5042 1.1836 2.5388	ें ह	2.0152 2.0155 1.6648 1.2209 .7795
	S		Ao Aos	
	#	.1155 1702 1315 1.1534 0165	& 8 -≠	.0020 .0031 .0144 1.1808
₹	n		ξς 1 10	0204 0093 1.6623 .0144 007
	Ø	4599 1.1291 3354 354 1.5470	cu	2.0014 2.0076 .0033 .0031
	н	1.8877 .1779 6255 2126 .3222	ਜ	2.03/2 0014 0204 .0000
		навчи		NEWBH
	ы	1.05 1.1.1 1.01 1.01 1.02 1.03	8	.0004 .0003 .0039 .0037 .9970
	τ. Ιν	24 48 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5	æ	
٠ دع	ก	73. 81. 82. 83. 84. 87. 87. 87.	8 P	. 0002 1.0042 .0009 .0009 .0009 1.0069
رة ا	C4	1.03	, A G	5000 5000 5000 5000 5000 5000 5000 500
. & ≺	н	2084485	н	1.0007 0005 0055 0055 0056 0056

**មពស**ងស**មិ**ភ

น<sub>์</sub>ผมน<sub>า</sub>นนิม

enter des la constanta de la constanta del la constanta del la constanta de la constanta de la constanta de la constanta del la c

<u>;</u>,

	TO LATERT VECTORS
	8
Eble 44	TAKE FOR RODATIONS
ä	FOR
	TABLE
	FASCILITATING

			•	
ပ	ļ	รูลูรุลูรุลูลูลูล	<i>ខំ</i> ន់ខំនុំនំងំនំងំ	
	욁	,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,	44444888888 624448888	QH A
م	From	\$	88882828888888888888888888888888888888	10 or more epproximately
υ		ម្ដង់មួយមួយមួយ	ૹ૽ૹ૽ <i>ૡ૽</i> ૹ૽ૹ૽ૡ૽ૹ૽ઌ૽ૹ૽	of 10 c
	욁	99955555999999999999999999999999999999	,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,	A so
م	Tron	ઌઌઌઌઌ ઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌઌ	8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8	
υ		ૄૹ૽ઌ૽ૹ૽ઌ૽૱ ૹ૽ઌ૽ૹ૽ઌ૽ૹ૽ઌ૽	ૄૻૢૢૢૢૢૢ૽ૢ૽ૢ૽ૢૢૢૢૢૢ૽ૢૢૢૢૢૢૢૢૢ૽૽ૢૢૢૢૢૢૢૢૢ	•
	욁			
م	Pron	64444444444444444444444444444444444444	44444444444444444444444444444444444444	
υ		\$\$£\$\$\$\$\$\$\$	<i>ૄ૾ૢૢૢૢૢ૽ૡ૽ૢ</i> ૢૢૢૢૢ૽ૡ૽ૹ૽ૡ૽ૹ૽	( <del>e</del>
م	क्ष म्ल	<i>ᡶᡱᡢᡳ</i> ᠌ᢧᠣ <i>ᡴ</i> ᡘ᠔ᠮ ᠼᡎᢋ <i>ᡳᢅᢧ</i> ᡢᡢᡊᡩᡩ	568 99 89 89 89 89 89 99 99 99 99 99 99 99	(Np244
	昌	• • • • • • • • •	· · · · · · · · · · · · · · · · · · ·	- <b>3</b> N 11 α
υ		88828848848	ស <u>៉ូ</u> ន្ល់ <u>ទុំ</u> ន្តំនុំទំនុំទំនុំទំនុំ	•
م	띩	98999999	4004888444	
	For	8 क्षेत्र के <b>व</b> ं प्रेम प्रस्	หู่ กู่ คู่ ชู	

#### 5. Notes on Matrix Computations

The computing procedures in the foregoing sections have been stated in terms of a few standard matrix computations described in this section. Familiarity with the following matrix computational methods will be of distinct assistance in understanding the directions in the first four sections.

#### a. Definitions:

1)	A row	of	numbers	is	called	8.	row vector.
	Examp	ple	:				

3 2 6 8 5

2) A column of numbers is called a column vector.

## Example:

3

2

6

8

5

5) A rectangular table of numbers is called a matrix.

#### Example:

7 2 1 4

3 4 2 5

8 1 7 5

4) A square matrix with entries in the diagonal from upper left to lower right and zeros elsewhere is called a diagonal matrix.

## Example:

5000

0 4 0

0 0 8 0

0 0 0 9

5) A single letter may be used to designate an entire vector or matrix.

### Examples:

Note that the following examples are different than the preceeding examples. The terms vector and matrix can be applied to any set of numbers arranged as a row, a column, and a table.

1) Row vector A equals:

7 5 9

2) Column vector B equals:

1 6 h

3) Matrix C equals:

58728

6) The transpose of a matrix is the matrix with the rows of the original matrix written as columns (or original columns written as rows which gives the same result). The transpose is designated by the letter for the original matrix primed.

### Example:

The transpose of matrix C is C: and is:

5 3 2 8 7 8

b. Multiplication of vectors.

Consider the following two row vectors

 Vector A
 7
 5
 9

 Vector E
 8
 3
 2

Multiply the first number of vector  $\underline{\mathbf{A}}$  by the first number of vector  $\mathbf{R}$ :

$$7 \times 8 = 56$$

Also find the product of the second entries in the two vectors:

Similarly find the products of each piar of corresponding entries in the two vectors. The third terms in the example give:

$$9 \times 2 \times 18$$

Sum the products:

$$56 + 15 + 18 = 89$$
.

This sum is the result from multiplying the two vectors. Thus:

Two vectors are multiplied by summing products of corresponding entries in the two vectors.

Consider a second e. Type; multiplication of row vector  $\underline{\mathbf{A}}$  and column vector  $\underline{\mathbf{B}}$ .

These two vectors are multiplied as follows:

$$7 \times 1 = 7$$
  
 $5 \times 6 = 30$   
 $9 \times 4 = 36$   
Total  $73$ 

The result of multiplying the two vectors is the number 73.

Note: Two rectors must have the same number of entries if they are to be multiplied.

#### c. Multiplication of Matrices:

Consider the following set of three matrices.

	Matrix C		Matrix F					Total	
	5	8			4	6	3	5	18
	3	7			1	9	2	7	19
	2	_ გ							
Total	10	23							

		Produ	ct Matr	ix CF	Total
	28	102	31	81	242
	19	81	23	64	187
	16	84	22	66	188
Total	63	267	76	211	617

Consider the first row of matrix C as a row vector. Consider the first column of matrix F as a column vector.

Multiplication of these vectors yields:

$$5 \times 4 = 20$$
  
 $8 \times 1 = 8$   
Total 28

The local 28, is recorded in the first row and first column of the product matrix CF.

Hultiplication of the second row of matrix  $\underline{C}$  by the first column of matrix  $\underline{F}$  yields:

$$3 \times 4 = 12$$
 $7 \times 1 = 7$ 
Total 19

19 is recorded in the second row and first column of the product matrix  $\underline{CF}$ .

Similarly, the 16 in the third row and first column of product matrix <u>CF</u> results from multiplying the third row of matrix <u>C</u> by the first column of matrix <u>F</u>.

When the row of column totals of matrix <u>C</u> is multiplied by the first column of matrix <u>F</u> the result is the total of the first column of the product matrix <u>CF</u>.

$$10 \times 4 = 40$$

$$25 \times 1 = 25$$

$$+01 = 1$$

This computation of the total of the column of the product matrix is an efficient check on the computation of the entries in the column.

The second column of the product matrix <u>CF</u> is obtained by multiplying the rows of matrix <u>C</u> by the second column of the matrix <u>F</u>.

First row,	second column	Third row,	second column
	5 x 6 = 30 8 x 9 = 72 Total 102 second column 3 x 6 = 18 7 x 9 = 63 Total 81	·	2 x 6 = 12 8 x 9 = 72 Total 84 second column 10 x 6 = 60 23 x 9 = 207 Total 207
	Total 81		

The third column of the product matrix C is obtained similarly by multiplying the rows of matrix C by the third column of matrix F. For the fourth column of the product matrix CF, the fourth column of matrix F is used. A check on the totals of the rows of the product matrix CF is obtained by multiplying the rows of matrix C by the column of totals of the rows of matrix F, for example:

5 x 18 = 90 8 x 19 = 152 Total 242

Multiplication of the row of totals of columns in matrix  $\underline{C}$  by the matr's  $\underline{F}$  column of totals of rows yields the grand total of all entries in the product matrix  $\underline{CF}$ .

10 x 18 = 180 23 x 19 = 437 Total 617

Two matrices are multiplied by multiplying as vectors each row of the first matrix by each column of the second matrix and recording the results in a product matrix with a row for each row of the first matrix and a column for each column of the second matrix.

Note that, for two matrices to be multiplied, the second matrix must have the same number of rows as the first matrix has number of columns.

Note, also, that the order of the matrices makes a difference in the matrix product. With square matrices care must be taken to consider the matrices in the proper order. Multiplication of matrices in the wrong order will produce erroneous results.